Rationale for azithromycin in COVID-19: an overview of existing evidence

Iwein Gyselinck, Wim Janssens, Peter Verhamme, Robin Vos

ABSTRACT
Azithromycin has rapidly been adopted as a repurposed drug for the treatment of COVID-19, despite the lack of high-quality evidence. In this review, we critically appraise the current pharmacological, preclinical and clinical data of azithromycin for treating COVID-19. Interest in azithromycin has been fuelled by favourable treatment outcomes in other viral pneumonias, a documented antiviral effect on SARS-CoV-2 in vitro and uncontrolled case series early in the pandemic. Its antiviral effects presumably result from interfering with receptor mediated binding, viral lysosomal escape, intracellular cell-signalling pathways and enhancing type I and III interferon expression. Its immunomodulatory effects may mitigate excessive inflammation and benefit tissue repair. Currently, in vivo reports on azithromycin in COVID-19 are conflicting and do not endorse its widespread use outside of clinical trials. They are, however, mostly retrospective and therefore inherently biased. The effect size of azithromycin may depend on when it is started. Also, extended follow-up is needed to assess benefits in the recovery phase. Safety data warrant monitoring of drug–drug interactions and subsequent cardiac adverse events, especially with hydroxychloroquine. More prospective data of large randomised controlled studies are expected and much-needed. Uniform reporting of results should be strongly encouraged to facilitate data pooling with the many ongoing initiatives.

INTRODUCTION
Since December 2019, the pandemic spread of the new virus SARS-CoV2 has affected over 50 million people. COVID-19—the disease caused by this virus—has killed over one million people in these past few months. Tremendous progress has already been made in the understanding of the disease. Still, only a few interventions have proven clinically beneficial and, besides thromboprophylaxis, these are mostly reserved for selected patients with an advanced disease stage. Their impact on the global disease burden, therefore, remains limited.

A high initial viral load and occurrence of a disproportional inflammatory response thereafter, the so-called cytokine storm, relate to adverse outcomes and are potentially modifiable. Hence, they are the target of most currently considered therapeutic strategies. Interference with the viral cycle is pursued through (1) inhibition of viral cell entry with TMPRSS2 inhibiting molecules such as camostat or aprotinin, (2) inhibition of viral lysosome escape with molecules as hydroxychloroquine, (3) antiretroviral drugs such as lopinavir/ritonavir that interfere with post-translational processing through the main protease and (4) inhibition of viral RNA-dependent RNA-polymerase with remdesivir or favipiravir. The excessive host’s inflammatory response is mitigated by (1) broad-spectrum molecules as dexamethasone or (2) targeted drugs as tocilizumab (anti-interleukin-6 (IL-6)), anakinra (anti-IL-1) or baricitinib (Janus kinase inhibitor). Finally, anticoagulants are effectively used to counter the inflammation-induced hypercoagulative state. Overall, time pressure has sparked a special interest into the repurposing of marketed or late stage molecules for COVID-19, parallel to the development of new and more selective drugs.

A repurposing drug candidate of special interest is azithromycin. Azithromycin is a macrolide antibiotic with a broad gram-positive and gram-negative spectrum. Moreover, it has well-documented anti-inflammatory and immunoregulatory effects, through modulation of both the innate and adaptive immune response. These are effective to treat chronic inflammatory disorders such as diffuse bronchiolitis, post-transplant bronchiolitis, non-eosinophilic asthma or rosacea. Azithromycin has also been associated with improved outcome in other viral pneumonias, such as influenza and rhinovirus, and in patients with acute lung injury admitted to the Intensive Care Unit (ICU). This has in some centres led to the early adoption of azithromycin in routine COVID-19 care.
further fuelled by reports of in vitro activity against SARS-CoV-2, and a suggested benefit in non-controlled case series early on in the SARS-CoV-2 pandemic. While more data of randomised controlled studies are eagerly awaited, we comprehensively review the rationale of its use against SARS-CoV-2, its window of opportunity and its possible limitations.

**PATHOPHYSIOLOGY OF COVID-19**

**Normal antiviral response**

SARS-CoV-2 is a positive-sense single stranded enveloped RNA β-coronavirus that spreads through aerosols, droplets, respiratory secretions and direct contact. One can distinguish different disease stages (figure 1). (A) After transmission, SARS-CoV-2 binds and enters respiratory epithelial cells through the ACE II (ACE2) receptor. The quick viral replication and high cytopathogenicity cause a strong release of danger signals, (B) Binding of these danger signals to specific pattern recognition receptors induces an innate antiviral immune response and clinical disease becomes apparent, (C) In the following days an adaptive immune response is gradually mounted, comprising a T-helpers-1 (Th1) and often also a Th2 activation. In the latter case, anti-SARS-CoV-2 IIgM and IgG antibodies appear and their levels correlate with disease severity. Assuming the patient is able to overcome the infection, a convalescent phase commences and (D) Inflammatory markers decrease and, in most patients, pulmonary infiltrates slowly wane.

**Excessive inflammatory response: cytokine storm**

Severe COVID-19 is characterised by a disproportional inflammatory response. This has been attributed to multiple traits of SARS-CoV-2, some in analogy with SARS-CoV and Middle East respiratory syndrome (MERS) (figure 2).

First, SARS-CoV-2 interferes with the innate antiviral immune response. Normally, two different antiviral pathways are activated. On the one hand, interferon (IFN) regulatory factors increase transcription of mainly type I and type III IFN, which stimulate natural killer cells and CD8 +cytotoxic T-lymphocytes. On the other hand, nuclear factor-κB (NF-κB) signalling promotes monocyte activation and their differentiation into M1 macrophages. These release proinflammatory cytokines and promote inflammatory T-cell (T_h1/T_h17) activation. SARS-CoV-2 skews the innate response towards macrophage activation. It suppresses type I and III IFN-related gene transcription, thereby favouring NF-κB activation. This impairs the recruitment of cytotoxic effector T-lymphocytes and causes abundant cytokine release and inflammasome formation. In severely ill and ICU-admitted patients, macrophage-related cytokines IL-6, IL-10 and TNFα are indeed consistently elevated compared with milder cases.

Second, excessive release of cytokines increases the expression of T-cell exhaustion markers, like programmed death 1 and T cell immunoglobulin and mucin domain-containing protein 3 (Tim-3). Together with IL-6 induced lymphocyte apoptosis and necrosis, T-cell exhaustion further dampens the cellular immune response. Lymphopenia is frequent and correlates with inflammation markers and disease severity.

Third, binding of SARS-CoV-2 to ACE2 receptors, and their subsequent internalisation, reduces ACE2-mediated angiotensin II breakdown. The increased angiotensin II levels enhance the inflammatory response, activate endothelial cells and locally increase
vascular permeability. This promotes coagulation by activation of the kallikrein-bradykinin system. A hypercoagulable state importantly contributes to COVID-19 morbidity and mortality. Finally, the excessive inflammation causes concern of pulmonary fibrosis as a possible late COVID-19 complication. In analogy with SARS and MERS, fibrotic changes have indeed been recognised in autopsy studies and may be associated with increased expression of tumour growth factor beta (TGF-β) and connective tissue growth factor. At this stage, it is still unclear who will recover, and who will proceed to uncontrolled cellular proliferation and persistent fibrotic remodelling.

**Rationale for Azithromycin Use in COVID-19**

**Pharmacological profile**

Azithromycin is a 15-membered-ring macrolide of the azalide class. It is safe and, besides mild gastrointestinal side effects, usually well tolerated. QT-prolongation and cardiotoxicity are a concern, especially when combined with other QT-prolonging drugs. However, while clearly demonstrated for the 14-membered-ring macrolides such as erythromycin and clarithromycin, few reports relate azithromycin to cardiac adverse events. As opposed to 14-membered-ring molecules, azithromycin is not metabolised by cytochrome P450 (CYP450), which accounts for a more favourable drug-drug interaction profile.

Azithromycin is rapidly absorbed after oral intake and has a long half-life. Its large volume of distribution is due to a high intracellular accumulation, with tissue concentrations up to a 100-fold higher than in plasma. The uptake is particularly high in leukocytes, but also in epithelial cells and fibroblasts. Intracellularly, it
has an affinity for acidic organelles such as lysosomes. Azithromycin also crosses the blood–brain barrier and concentrates in central nervous system tissue. This is noteworthy as there is increasing awareness of neurological complications of COVID-19, due to infiltration and activation of residing inflammatory cells and possibly direct viral neurotropism.

In vitro data on the inhibitory concentrations of azithromycin on SARS-CoV-2 and other viruses have recently been summarised elsewhere. However, these data are scarcely replicated and far from an in vivo pharmacokinetic-pharmacodynamic target. On the other hand, azithromycin accumulation in leukocytes ensures effective delivery to sites of infection and inflammation. In vivo lung tissue homogenates reach concentrations well above the reported 90% effective concentration after 3 days of oral therapy with 500 mg azithromycin. Similar regimens are approved and long used to treat bacterial gastroenteritis and respiratory tract infections. Slightly longer treatment durations of 5 up to 8 days were evaluated in cohorts studies assessing the effect of azithromycin in hospitalised patients with influenza or ICU patients with acute lung injury.

**Antiviral effects**

Azithromycin has direct and indirect antiviral activity in bronchial epithelial cells and other host cells. In addition to SARS-CoV-2, this has also been shown for influenza, rhinovirus, dengue, ebolavirus, parainfluenza virus, zika virus and enterovirus.

There are multiple mechanisms for azithromycin’s antiviral effect. For host-cell entry, the prerequisite binding of the SARS-CoV-2 viral spike protein to ACE2 has been repeatedly described. Virtualised mechanical modelling techniques demonstrated that azithromycin may interfere due to its affinity with the binding interaction point of the spike protein and ACE2. Also, azithromycin may competitively inhibit a viral cofactor binding site due to its striking molecular similarity with GM1, a host-cell ganglioside that binds the ganglioside binding domain of the spike protein. Further experimental work is needed to confirm these possible modes of action. After receptor binding, the virus enters host cells either through membrane fusion, or through receptor mediated endocytosis. In the second route, endosome acidification facilitates viral escape and subsequent release of the nucleocapsid. Azithromycin interferes at this level, as it is a weak base that accumulates intracellularly and inside endosomes.

During the remainder of the viral cycle, viruses are known to hijack intracellular antiapoptotic signalling pathways to promote their survival and replication. As an example, blocking the PI3K/AKT/mTOR-pathway decreases MERS-CoV replication in vitro. Rapamycin (sirolimus) is a known mammalian target of rapamycin (mTOR) -inhibiting macrolide, but azithromycin has also shown to interfere with mTOR-signalling, although to a lesser extent. It remains unclear if and how this affects SARS-CoV-2 replication.

Furthermore, azithromycin also has indirect antiviral effects. It induces intracellular mRNA expression of antiviral genes, IFN-stimulated genes and IFN production in infected host cells. This may enhance the cellular antiviral response mediated by the IFN pathway and help to retain balance in the early innate immune response.

**Anti-inflammatory effect and modulation of macrophage action**

Azithromycin has well-documented immunomodulatory properties, that may affect the disease course of COVID-19.

First, in in vitro models with respiratory epithelial cells azithromycin decreases mucus production and increases epithelial barrier thickness. It also reduces matrix metalloproteinase (MMP) activity after challenge with bacterial lipopolysaccharides. This reduces inflammatory signalling, and helps to remain cell integrity and epithelial barrier function. These experiments have not been replicated with viral antigens. However, the related macrolide clarithromycin has shown to decrease lung and serum MMP-9 levels and vascular hyperpermeability due to influenza A infection in mouse models.

Second, azithromycin is a potent modulator of monocyte and macrophage cytokine responses. It may balance the immune answer in COVID-19 by suppressing NF-kB signalling and reducing release of classical M1 activated macrophage differentiation markers IL-8, IL-6, TNFα and granulocyte-macrophage colony-stimulating factor. Azithromycin promotes polarisation of macrophages from a M1 to an M2 phenotype, thereby augmenting their phagocytic capacity.

Third, azithromycin also modulates Th2-cell and B-cell responses. For example, it reduces the serum titre of specific IgG1-antibodies after vaccination with pneumococcal conjugate vaccine in mice. It is yet unclear how the antibody response contributes to the pathophysiology of COVID-19. Late neutralising antibodies seem to be protective. However, early IgG-response has been associated with more severe disease, possibly due to antibody-dependent enhancement.

Fourthly, azithromycin attenuates neutrophil function. It downregulates chemoattractants and adhesion molecules in activated vascular endothelial cells, reduces neutrophil activation and constrains the release of neutrophil extracellular traps (NET). Neutrophilia and NETosis contribute to hyperinflammation and hypercoagulability in severe COVID-19, but may be secondary to other processes like bacterial coinfection.

Finally, azithromycin attenuates TGF-β-induced myofibroblast differentiation, fibroblast collagen secretion and extracellular matrix remodelling. This occurs through a decrease of both MMP production and vascular endothelial growth factor release. Eventually, this limits the
damaging effects of inflammation, fibrosis formation and vascular remodelling.

**Prophylaxis against bacterial superinfection**
The reported rate of antibiotic prescription in COVID-19 patients, especially in-hospital, is very high. Driving forces are the sometimes-difficult differential diagnosis with atypical pneumonias and fear of bacterial superinfection. Early bacterial coinfection has indeed been a well-known source of morbidity and mortality in historic influenza pandemics. In COVID-19, however, pooled data suggest a much lower risk of bacterial co-infection, and do not support routine administration of antibiotics.

Even though azithromycin may improve outcomes in the limited cases of superinfection, antibacterial prophylaxis is no ground argument for its systematic use, and must be weighed against the risk of bacterial resistance.

**In vivo data**

**Non-COVID-19**

Azithromycin is an established treatment modality in several chronic inflammatory respiratory diseases. Different clinical trials have proven its efficacy in chronic obstructive pulmonary disease, bronchiectasis, asthma and lung transplantation. While undeniable proof of azithromycin’s immunomodulatory potential, it is unsure if this can also be exploited in the acute setting.

Before COVID-19, the anti-inflammatory and antiviral effects of azithromycin have been clinically demonstrated in several viral pneumonias and in acute respiratory distress syndrome (ARDS). In a retrospective cohort evaluation of hospitalised patients with moderate or severe ARDS treated with azithromycin or not, azithromycin was associated with a significant improvement in 90-day survival rate and a shorter time to successful discontinuation of mechanical ventilation. Also, azithromycin-use was associated with decreased 60-day mortality and shorter time of ventilator dependency in patients with sepsis-associated ARDS. For the treatment of influenza, combination therapy of oseltamivir-azithromycin compared with oseltamivir alone showed improved clinical outcomes in a retrospective cohort and a faster decline of inflammatory parameters in a randomised controlled trial.

On the other hand, a tendency towards lower ICU mortality, lower 90-day mortality and shorter hospital stay did not achieve statistical significance in a cohort study on the use of macrolides (of which 71.3% was azithromycin) in critically ill patients with MERS. Possibly, the higher risk of coinfection in influenza, especially with influenza A, may contribute to the larger effect size.

**COVID-19**

The positive reports on azithromycin in other respiratory viral diseases have prompted the rapid initiation of interventional trials to evaluate its efficacy in COVID-19. At the time of writing, 121 trials with azithromycin are listed in clinicaltrials.gov. At the start of the pandemic, however, following the example of early non-randomised series of a French group in Marseille, azithromycin has most often been prescribed as an adjuvant to hydroxychloroquine. The use of hydroxychloroquine is now largely abandoned and few published studies have assessed azithromycin alone. The reported effects of azithromycin are thus often derived from patients treated with hydroxychloroquine-azithromycin combination versus hydroxychloroquine alone. Table 1 gives an overview of currently published peer-reviewed studies in the MEDLINE database, in which the effect of azithromycin is assessed. Studies only comparing combination regimens versus standard of care were not considered (eg, hydroxychloroquine and azithromycin vs neither therapy), as no inference about the individual treatment effect of azithromycin could be deduced (see online supplemental material for detailed description of the individual studies and study selection).

Studies that assess azithromycin monotherapy versus standard of care in hospitalised patients report a wide effect range, from a decreased adjusted OR for mortality of 0.60 (95% CI 0.42 to 0.85) in the retrospective cohort of Albani et al to a non-significantly increased adjusted OR of 1.30 (95% CI 0.65 to 2.64) in Kudery et al. Even more heterogeneity is seen in studies that assess the addition of azithromycin to hydroxychloroquine, with a survival benefit (adjusted HR of 0.294; 95% CI 0.218 to 0.396) seen by Arshad et al opposed to a significantly increased 30-day mortality (adjusted OR 2.93; 95% CI 1.79 to 4.79) reported again by Kudery et al. In an outpatient setting, Guérin et al reported a significant reduction in the mean time to clinical recovery with azithromycin (12.9 days with azithromycin vs 25.8 days without; p<0.0001). A significant difference in hospitalisation risk was, however, not withheld by Szente et al (adjusted OR for azithromycin-containing vs no-azithromycin-containing regimens 0.93; 95% CI 0.72 to 1.90).

The increased mortality reported for hydroxychloroquine-azithromycin combination by Kudery et al together with increased incidence of adverse events of this regimen in Rosenberg et al and the randomised controlled trial of Cavalcanti et al strengthen the concerns about QT-prolonging drug-drug interactions. Importantly, no studies reported a significantly increased risk of adverse outcomes with azithromycin monotherapy. Cavalcanti et al did not assess efficacy of azithromycin monotherapy, but found no increased adverse events in this treatment group, whereas QTc prolongation and increased transaminases were seen in the hydroxychloroquine containing regimens. Similarly, Rosenberg et al reported an increased incidence of cardiac arrest with hydroxychloroquine and azithromycin coadministration (adjusted OR, 2.13; 95% CI 1.12 to 4.05) and when comparing hydroxychloroquine monotherapy with azithromycin monotherapy (adjusted OR, 2.97; 95% CI 1.56 to 5.64) but not for azithromycin vs neither drug (adjusted OR, 0.64; 95% CI 0.27 to 1.56).

The interpretation of these heterogeneous results is troublesome in many ways. First, estimations of
azithromycin’s individual treatment effect from combination regimens with hydroxychloroquine may be unsound. Drug-drug interactions may increase short-term mortality and follow-up is often short to assess any long-term azithromycin benefits (eg, progression to fibrosis). Second, most of the studies are retrospective. State-of-the-art statistical corrections like propensity score weighting are used in nearly half of the retrospective studies, but the propensities are often calculated on baseline patient characteristics like age, sex, comorbidities, obesity, while factors that have now been clearly associated with combination treatment vs neither drug. Moreover, initiation of any form of treatment has been influenced by various factors other than baseline characteristics and disease severity, such as drug availability, do-not-resuscitate orders and changing local policies. Third, the difference in techniques to adjust for confounders, but also the difference in primary outcomes (clinical improvement, mortality, hypoxia, hospitalisation risk), outcome measures (comparing odds vs time-to-event and survival analyses), target populations (mild vs severe, outpatients vs hospitalised patients) and follow-up times (in hospital mortality, 30-day mortality) all contribute to the heterogeneity and hinder data pooling for meta-analyses. We summarised the published meta-analyses that pooled azithromycin containing regimens (see online supplemental table A). They confirm the increased mortality risk in hydroxychloroquine–azithromycin cotreated patients. However, as they are largely based on the sometimes heavily biased data of the studies discussed above, one might still doubt a causal inference. The data of azithromycin monotherapy have not been pooled, and of the three meta-analyses that directly compared hydroxychloroquine with azithromycin versus hydroxychloroquine alone, only Das et al.77 found a significantly increased mortality with the addition of azithromycin. Interestingly, not cardiac adverse events but rather the development of severe disease was an outcome associated with the addition of azithromycin to hydroxychloroquine. As there is no mechanistic rationale to expect disease worsening with azithromycin, this may as well signal residual indication bias.

Overall, the limited and low-quality evidence does not endorse azithromycin’s widespread use in the treatment of COVID-19. On the other hand, monotherapy is safe and therefore justifiable in a clinical trial setting. The data at least urges close monitoring when combined with other QT-prolonging drugs like hydroxychloroquine, or when other risk factors for long QT exist. A risk mitigation strategy such as applying strict ECG criteria to initiate (eg, only if QTc <450) and halt (eg, if QTc exceeds 500 ms
or increases>60 ms since start of treatment) azithromycin may be warranted.78-80

**DISCUSSION**

The use of azithromycin in COVID-19 is mechanistically well grounded and indirectly supported by prior experiences with other viral pneumonias, chronic pulmonary diseases and inflammatory disorders. Yet, the empirical practice of azithromycin treatment for COVID-19 has not been substantiated by good quality clinical data. Despite—maybe even because of—the limitations, a critical appraisal of the currently available evidence is valuable. It should contextualise the results of ongoing trials and could improve the set-up of future trials.

First, most interventions have an optimal time window. From a mechanistic point of view, initiation of azithromycin before or during the early inflammatory phase is more sensible. At that early stage, an antiviral effect could still be relevant. It remains unclear, however, if azithromycin significantly inhibits viral replication in vivo. Better supported by the data in this review are the immunomodulatory effects of azithromycin on early inflammatory pathways that are key in the progression to severe COVID-19. They are supposed to balance the adaptive immune response, stimulate cellular immunity and avoid a subsequent cytokine storm. Results of large randomised controlled trials for hospitalised patients (eg, RECOVERY)81 are soon expected. However, a significant share of hospitalised patients may already be beyond this window. The primary care setting may be more suited to evaluate early interventions. Compared with the hospital though, this is a much less controlled environment, which makes retrospective data collection very challenging.

A few studies are published, and the positive signals of Guérin et al82 and Esper et al83 (preprint article, not included in table 1) are contradicted by Sznite Fonseca et al.74 At least, with only a short follow-up time needed to assess the risk of hospital admission, prospective data in this context (eg, ATOMIC2, ACTION)83 84 should soon be able to provide more clarity.

Second, despite the pleiotropic effects of azithromycin, it is certainly not the most potent molecule. Targeted antiviral drugs will likely have a more robust effect on the viral load. However, experience with influenza has taught us to start antivirals as soon as possible after host infection.85 Likewise, the anti-inflammatory effects of targeted anti-IL1, anti-IL6 or steroids are stronger, though probably only warranted when clear signs of hyperinflammation are present.86 If anything, one should not expect azithromycin to be put forward as ‘the standard treatment’, but rather as a part of a multimodal approach of antiviral, antithrombotic, anti-inflammatory and—in selected cases—antibiotic drugs, depending on the patient’s presentation, immune status and disease stage.

Lastly, it is important to consider treatment effects that surpass acute pulmonary inflammation. Azithromycin has antifibrotic properties and crosses the blood–brain barrier. Possible morbidity of sequellar fibrotic lung disease and of prolonged neurological complaints extends well beyond the acute phase, and attenuating this later phase will significantly impact quality adjusted life years of COVID-19 patients. A comprehensive clinical trial assessment with extended follow-up is, therefore, crucial to confirm or exclude the hypothetical benefits of azithromycin in COVID-19.

In conclusion, its favourable safety profile, affordability and pleiotropic mechanisms have raised a large interest in azithromycin to treat COVID-19. Its effect on the early inflammatory phase is best supported by the current evidence, which is typically when the first symptoms arise and a patient contacts his caretaker. Before starting azithromycin, a comprehensive assessment for drug–drug interactions and cardiovascular risk factors is prerequisite, especially when use in the first line is advocated. Beyond that, the current data remain equivocal. Due to the scale of the current pandemic, however, even a small treatment effect could mean a significant absolute reduction in COVID-19-related morbidity and mortality. Moreover, we have currently no idea on how a second primary infection will be eradicated by the hosts’ primed immune system. Beneficial modes of action should not be discarded based on short-term results obtained during the first wave of hospital admissions. In the next months, results of adequately performed randomised trials will provide better insight into the true role of azithromycin and other repurposed drugs in this historic pandemic. Still, as the field of intervention studies in COVID-19 is currently highly scattered, large coordinated international initiatives will be needed to pool aggregated and individual patient data to come to optimal conclusions.87

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References


46. Fantini J, Chahinian H, Yahia N. Synergic antiviral effect of hydroxychloroquine and azithromycin in combination against variety of respiratory viruses.


Supplementary figure A: study selection flowchart

Medline search: 535
Search term: ("covid-19" OR "SARS-COV-2" OR "Severe acute respiratory distress syndrome coronavirus 2") AND "azithromycin"

Other sources: 2 [1, 2]

Preclinical, hypothetical, narrative review, opinion, article not in English: 481

Case reports, no clinical endpoint, small case series (<50), no control group: 30 [3-32]

Articles included:
- Original research: 20 [1,2, 33-50]
- Meta-analyses: 6 [51-56]
## Supplementary table A: studies assessing azithromycin (monotherapy and combination therapies)

<table>
<thead>
<tr>
<th>Reference</th>
<th>Study type and setting</th>
<th>Treatment/intervention</th>
<th>Outcome</th>
<th>Results</th>
<th>Safety</th>
<th>Limitations, remarks</th>
</tr>
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<tbody>
<tr>
<td>Kuderer et al.</td>
<td>Multicentric, retrospective cohort study, Cancer patients with confirmed or probable diagnosis, in- and outpatients</td>
<td>HQ alone (n = 89) AZ alone (n = 93) HQ + AZ (n = 181) Neither (n = 486)</td>
<td>30-day mortality</td>
<td>Multivariable adjusted odds ratios for all-cause mortality:</td>
<td>Not reported</td>
<td>- Adjusted for baseline patient characteristics, but not for disease severity</td>
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<tr>
<td>May 2020</td>
<td></td>
<td>Dosing: Not reported</td>
<td></td>
<td>- HQ vs neither: OR, 1.06; 95% CI, 0.51 - 2.20</td>
<td></td>
<td>- Secondary endpoint of severe illness (composite of death, hospital admission, ICU admission) was associated with both AZ or HQ + AZ, for which indication bias by disease severity is a more plausible explanation than worsening with association of azithromycin</td>
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<td>[33]</td>
<td></td>
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<td></td>
<td>- AZ vs neither: OR, 1.30; 95% CI 0.65 - 2.64</td>
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<td></td>
<td>- AZ + HQ vs neither: OR, 2.93; 95% CI, 1.79 – 4.79</td>
<td></td>
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</tr>
<tr>
<td>Geleris et al.</td>
<td>Monocentric, retrospective cohort study, USA Hospitalized patients with confirmed infection</td>
<td>HQ + AZ (n = 486) HQ alone (n = 325) AZ alone (n = 127) Other (n = 438)</td>
<td>Time from study baseline to intubation or death (for patients who died after intubation, the timing of the primary end point was defined as the time of intubation)</td>
<td>Multivariable Cox model with inverse probability weighting according to propensity score for composite endpoint:</td>
<td>Not reported</td>
<td>- Data extracted from clinical data warehouse; no data were manually extracted from electronic medical records</td>
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<tr>
<td>June 2020</td>
<td></td>
<td>Dosing: HQ: 600mg bid day 1, 400mg od day 2-5 AZ: 500mg od day 1, 250mg od day 2-5</td>
<td></td>
<td>- no significant association between treatment with azithromycin and the composite end point (hazard ratio, 1.03; 95% CI, 0.81 to 1.31).</td>
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<tr>
<td>Rosenberg et al.</td>
<td>Multicentric, retrospective cohort study, USA Hospitalized patients with confirmed infection</td>
<td>HQ alone (n = 271) AZ alone (n = 211) HQ + AZ (n = 735) neither (223)</td>
<td>In hospital mortality</td>
<td>Adjusted Cox regression hazard ratio for mortality:</td>
<td>Cardiac arrest more likely in HQ + AZ but not in either AZ alone or HQ alone</td>
<td>- Cardiac arrest more likely in HQ + AZ but not in either AZ alone or HQ alone</td>
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<tr>
<td>June 2020</td>
<td></td>
<td>Dosing: Different regimens</td>
<td></td>
<td>- HQ + AZ vs neither: HR, 1.35; 95% CI, 0.76 - 2.40; p = 0.31</td>
<td></td>
<td>- Adverse events recorded at any point during hospitalization, potentially before drug initiation</td>
</tr>
<tr>
<td>[34]</td>
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<td></td>
<td>- HQ alone vs neither: HR, 1.08; 95% CI, 0.63 - 1.85; p = 0.79</td>
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<td>- Some potential confounders (e.g. inflammatory markers) not available for multivariate analysis</td>
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<td>- AZ alone vs neither: HR, 0.56; 95% CI, 0.27 - 1.56; p = 0.14</td>
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<td>- Mortality endpoint was not adjusted for MV or CPAP</td>
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<tr>
<td>Authors</td>
<td>Type of Study</td>
<td>Hospitalized Patients with Confirmed Infection</td>
<td>Dosing</td>
<td>In-Hospital Mortality</td>
<td>Adjusted Cox Regression Hazard Ratio for Mortality</td>
<td>Safety, Tolerability, and Clinical Outcomes</td>
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<tr>
<td>Arshad et al.</td>
<td>Multicentric, retrospective cohort study, USA</td>
<td>HQ alone (n = 1202) AZ alone (n = 147) HQ + AZ (n = 783) Neither</td>
<td>HQ: 400mg bid day 1 and 200mg bid day 2-5 AZ: 500mg od day 1, 250 od day 2-5</td>
<td>HQ alone vs neither: HR, 0.340; 95% CI, 0.254 - 0.455; p&lt;0.001 AZ alone vs neither: HR, 1.050; 95% CI, 0.682 - 1.616; p = 0.825 HQ + AZ vs neither: HR, 0.294; 95% CI, 0.218 - 0.396; p&lt;0.001</td>
<td>No patient had documented torsade de pointes. Suggests no differences between treatment arms, although specific data not provided.</td>
<td>QTc increased above 500ms in 47 of 117 patients who had ECG follow up, of which 3 concomitantly used azithromycin. Only 1 patient developed non-sustained monomorphic VT and this was in the HQ alone group. No other ventricular tachycardia was reported.</td>
</tr>
<tr>
<td>Tanriverdi et al.</td>
<td>Monocentric, retrospective cohort study, Turkey</td>
<td>HQ alone (n=30) HQ + favipiravir (n = 9) HQ + lopinavir/ritonavir (n = 18)</td>
<td>HQ: 400mg bid day 1, 200mg bid day 2-10 AZ: 500mg od day 1, 250 od day 2-5</td>
<td>Duration of hospitalization: subgroup analysis of HQ + AZ vs HQ + other antibiotic: 6.68 days vs 8.16 days; p = 0.027 No difference in other outcomes including mortality, ICU admission</td>
<td>No unexpected arrhythmia or cardiac event observed.</td>
<td>Safety, tolerability and clinical outcomes (hypoxia, need for MV, mortality)</td>
</tr>
<tr>
<td>Satlin et al.</td>
<td>Multicentric, retrospective cohort study, USA</td>
<td>HQ alone (n = 132) HQ + AZ (n = 27)</td>
<td>HQ: 600mg bid day 1, 400mg od day 2-5 AZ: 500mg bid day 1, 250mg od day 2-5</td>
<td>Multivariate adjusted odds ratio for hypoxia improvement - HQ + ≥3 days of azithromycin vs HQ alone: multivariate adjusted OR, 0.99; 95% CI, 0.38 – 2.60; p not reported Multivariate adjusted odds ratio for mortality - HQ + ≥3 days of azithromycin vs HQ alone: OR, 1.14; 95% CI, 0.37–3.50; p not reported</td>
<td>More steroid use in treated patients (although corrected for in propensity matching, however no propensity matching was done for azithromycin effect) Immortal time bias Discrepancy between higher mortality and lower ICU stay in not-treated group may depend on patient characteristics not accounted for in multivariate analysis (e.g. no treatment because of palliative care)</td>
<td>Small sample size with low number of AZ patients Univariate prefiltering and small sample size may have excluded confounders from the multivariate model</td>
</tr>
</tbody>
</table>
### Cavalcanti et al. 
**July 2020 [38]**  
**Multicentric, open label randomized controlled trial, Brazil**  
Hospitalized patients with confirmed infection, mild to moderate disease  
<table>
<thead>
<tr>
<th>Group</th>
<th>Description</th>
<th>Dosing</th>
<th>Clinical status on day 15 on ordinal scale</th>
<th>Proportional odds of having a worse score at day 15:</th>
</tr>
</thead>
<tbody>
<tr>
<td>HQ alone (n = 159)</td>
<td>HQ + AZ (n = 172)</td>
<td>Dosing: HQ: 400mg bid for 7 days AZ: 500mg od for 7 days</td>
<td>Proportional odds of having a worse score at day 15:</td>
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</tbody>
</table>
| Neither (n = 173) | - HQ + AZ vs SOC: OR 0.99; 95% CI, 0.57 - 1.73; p=1.00  
- HQ alone vs SOC: OR, 1.21; 95% CI, 0.69 - 2.11; p=1.00  
- HQ + AZ vs HQ alone: OR, 0.82; 95% CI, 0.47 - 1.43; p=1.00 | - Safety population also included AZ alone patients.  
- More AE reported in HQ + AZ group or HQ alone group than in AZ alone group and neither group  
- Prolongation of QT and elevated liver enzymes were more common in HQ alone group or HQ + AZ group than in neither group (however more serial ECG follow up in treated patients) | - Point of estimate instead of cox regression |

### Guerin et al. 
**July 2020 [1]**  
**Prospective observational study in MDs and their relatives, France**  
Outpatients with flu-like symptoms with confirmed and suspected infection  
<table>
<thead>
<tr>
<th>Group</th>
<th>Description</th>
<th>Dosing:</th>
<th>Time to complete clinical recovery</th>
<th>Mean times to achieve clinical recovery</th>
</tr>
</thead>
</table>
| AZ alone (n = 34) | AZ + HQ (n = 20) | Dosing: AZ: 500mg od day 1, 250mg od day 2-5  
HQ: 600mg od for 7 to 10 days | In-hospital mortality | - Neither: 25.8 days  
- AZ: 12.9 days (p < 0.0001 for AZ vs neither)  
- AZ + HQ: 9.2 days (p < 0.0001 for AZ + HQ vs neither; p = 0.26 for AZ vs AZ + HQ) | - No serious adverse event nor cardiovascular events were reported in any treatment group (ECG done before initiation of HQ in all patients)  
- Gastrointestinal adverse events reported in treatment group  
- Similar results with Logrank analysis  
- Similar results in case-control analysis (3x19 patients matched for age, sex and body mass index) |
| Neither (n = 34) | Not reported | Not reported | Not reported | Not reported |

### Monforte et al. 
**July 2020 [39]**  
**Monocentric, retrospective cohort study, Milan**  
Hospitalized patients with confirmed infection  
<table>
<thead>
<tr>
<th>Group</th>
<th>Description</th>
<th>Dosing:</th>
<th>In-hospital mortality</th>
<th>Adjusted hazard ratio for in hospital mortality:</th>
</tr>
</thead>
</table>
| HQ alone (n = 197) | HQ + AZ (n = 94) | Dosing: not reported | In-hospital mortality | - HQ vs neither: HR, 0.66; 95% CI, 0.39 – 1.11; p = 0.118  
- HQ + AZ vs neither: HR, 0.44; 95%CI, 0.24 – 0.82; p = 0.009  
NB: treatment effectiveness was more substantial in less severe cases |
| Neither (n = 92), but 47 received other treatment (lopinavir, darunavir, steroids or other immunomodulatory drugs) | Not reported | Not reported | Not reported | Not reported |

### Ip et al. 
**August 2020 [40]**  
**Multicentric, retrospective cohort study, USA**  
Hospitalized patients with confirmed infection  
<table>
<thead>
<tr>
<th>Group</th>
<th>Description</th>
<th>Dosing:</th>
<th>30-day mortality</th>
<th>Propensity score stratification adjusted hazard ratio for 30-day mortality:</th>
</tr>
</thead>
</table>
| HQ alone (n = 441) | AZ alone (n = 256) | Dosing: heterogeneous | In-hospital mortality | - HQ alone vs no HQ: HR, 1.02; 95% CI, 0.83 – 1.27; p = 0.83  
- AZ alone vs no AZ: HR, 0.89; 95% CI, 0.72 – 1.10; p = 0.28  
- HQ + AZ vs neither: HR, 0.98; 95% CI, 0.75–1.28; p = 0.89 |
| HQ + AZ (n = 1473) | Neither (n = 342) | Not reported | Not reported | Not reported |

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<table>
<thead>
<tr>
<th>Study</th>
<th>Setting</th>
<th>Patients Description</th>
<th>Dosing</th>
<th>Outcomes</th>
<th>Factors Considered</th>
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</thead>
<tbody>
<tr>
<td>Magagnoli et al.</td>
<td>Multicentric, retrospective cohort study, USA</td>
<td>Hospitalized patients with confirmed infection</td>
<td>HQ alone (n = 198) HQ + AZ (n = 214) Neither (n = 395)</td>
<td>Mortality, use of MV</td>
<td>Propensity score adjusted (regression on propensity splines) hazard ratio for risk of death from any cause</td>
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<td>- HQ alone vs neither: HR, 1.83; 95% CI 1.16 - 2.89; p = 0.009</td>
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<td>- HQ + AZ vs neither: HR, 1.31; 95% CI 0.80 - 2.15; p = 0.28</td>
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<td>Propensity score adjusted hazard ratio for risk of mechanical ventilation</td>
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<td>- HQ alone vs neither: HR, 0.78–1.82; p = 0.42</td>
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<td>- HQ + AZ vs neither: HR, 1.09; 95% CI, 0.72–1.66; p=0.69</td>
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<td>Not reported</td>
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<tr>
<td>Sekhavati et al.</td>
<td>Monocentric, open label RCT, Teheran</td>
<td>Hospitalized patients with confirmed disease</td>
<td>AZ + HQ + LPV/R (n = 56) HQ + LPV/R (n = 55)</td>
<td>Vital signs, hypoxia, duration of hospitalisation, need for and length of intensive care unit admission, mortality rate and results of 30-day follow-up after discharge</td>
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<td>Duration of hospitalization</td>
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<td>- AZ group 4.61 days vs non-AZ group 5.96 days; p = 0.02</td>
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<td>Mean duration of ICU stay</td>
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<td>- AZ-group 5 days vs non-AZ group 4.43 days; p = 0.157</td>
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<td>NB: Also, better oxygenation at discharge for AZ-group</td>
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<td>No adverse events while using a risk scoring system to exclude patients at high risk for QT-prolongation</td>
</tr>
<tr>
<td>Albani et al.</td>
<td>Monocentric, retrospective cohort study, Italy</td>
<td>Hospitalized patients with confirmed infection</td>
<td>HQ alone (n = 211) AZ alone (n = 421) HQ + AZ (n = 166) Neither (n = 605)</td>
<td>In-hospital mortality</td>
<td>Overlap weighted propensity score adjusted odds ratio for in hospital mortality</td>
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<td>- AZ alone vs neither: OR, 0.60; 95% CI, 0.42–0.85</td>
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<td>- HQ alone vs neither OR, 0.76; 95% CI, 0.53–1.08</td>
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<td>- HQ + AZ vs neither: OR, 1.13; 95% CI, 0.77–1.69</td>
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<td>Factors that may have influenced treatment decisions (e.g. palliative care) or some measures for disease severity (lymphocytes, D-dimers) were not accounted for in multivariate regression</td>
</tr>
</tbody>
</table>

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<table>
<thead>
<tr>
<th>Author(s)</th>
<th>Study Type</th>
<th>Setting</th>
<th>Patients</th>
<th>Regimen Details</th>
<th>Outcomes</th>
<th>Study Details</th>
</tr>
</thead>
<tbody>
<tr>
<td>Rodriguez-Molinero et al.</td>
<td>Multicentric, retrospective cohort study, Spain</td>
<td>Hospitalized patients with confirmed infection</td>
<td>Regimen without AZ (n = 29) or with AZ (n = 29)</td>
<td>Matched subcohorts</td>
<td>O$_2$/FiO$_2$ at 48 hours after inclusion and length of hospital stay</td>
<td>Mean difference in O$_2$/FiO$_2$ at 48 hours of matched subcohorts: AZ vs no AZ: Δ O$_2$/FiO$_2$, 0.02%; 95% CI, -1.35% - 1.39%; p=0.980. Adjusted Cox regression hazard ratio for time to discharge in unmatched cohorts: AZ vs no AZ: HR, 1.45; 95% CI, 0.88-2.41; p=0.150. Insufficient events to draw conclusions. Small sample sizes, especially after matching.</td>
</tr>
<tr>
<td>Furtado et al.</td>
<td>Multicentric open label randomized controlled trial, Brazil</td>
<td>Hospitalized patients with confirmed infection, severe disease</td>
<td>HQ alone (n = 183) HQ + AZ (n = 214)</td>
<td>Dosing: HQ: 400mg bid 10 days AZ: 500mg od 10 days</td>
<td>Clinical status on day 15</td>
<td>Proportional odds of being in worse clinical category: AZ + HQ vs HQ: OR, 1.36; 95% CI, 0.94–1.97; p=0.11. Proportion of patients with any serious adverse event was 42% in the azithromycin group and 38% in the control group (p=0.35). Large estimated effect size limits power analysis.</td>
</tr>
<tr>
<td>Lauriola et al.</td>
<td>Monocentric, retrospective cohort study, Italy</td>
<td>Hospitalized patients with confirmed infection</td>
<td>HQ alone (n = 17) HQ + AZ (n = 297) neither (n = 63)</td>
<td>Dosing: HQ: 200mg bid 10 days AZ: 500mg od 10 days</td>
<td>In-hospital mortality</td>
<td>Adjusted Cox regression hazard ratio for in hospital mortality: HQ alone vs neither: HR 1.108; 95% CI, 0.536 – 2.293; p=0.782. HQ + AZ vs neither: HR, 0.265; 95% CI, 0.171-0.412; p&lt;0.001 No fatal arrhythmias have been observed during treatment. Small sample size, monocentric.</td>
</tr>
<tr>
<td>Ayerbe et al.</td>
<td>Multicentric, retrospective cohort study, Spain</td>
<td>Hospitalized patients with confirmed infection</td>
<td>HQ alone (n = 670) HQ + AZ (n = 1187) neither (n = 162)</td>
<td>Dosing: HQ: 400mg bid day 1, 200mg bid day 2-5 AZ: not reported</td>
<td>Mortality (over study window: March – April)</td>
<td>Multivariate logistic regression adjusted odds ratio for mortality for AZ + HQ vs HQ alone (3th of 4 tested models): main effect of AZ on mortality: OR, 0.53; 95% CI, 0.19-1.50; p = 0.233. No interaction effect between AZ and HQ on mortality: OR, 1.11; 95% CI, 0.38 – 3.29; p = 0.846. NB: HQ was associated with lower mortality: OR, 0.45; 95% CI, 0.30 – 0.68; p &lt; 0.001 Not reported.</td>
</tr>
<tr>
<td>Study</td>
<td>Design</td>
<td>Patients</td>
<td>Intervention</td>
<td>Dosing</td>
<td>Endpoints</td>
<td>Statistical Analysis</td>
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<tr>
<td>Lammers et al. Sep [48]</td>
<td>Multicentric, observational cohort study, The Netherlands</td>
<td>Hospitalized patients with confirmed infection or typical disease findings on CT, mild to moderate disease</td>
<td>HQ/CQ alone (n = 487) HQ/CQ + AZ (n = 131) AZ alone (n = 367) neither (n = 367)</td>
<td>Dosing: HQ: 400mg bid day 1, 200mg bid day 2-5 AZ: not reported</td>
<td>Death and ICU admission (composite endpoint)</td>
<td>Logrank test shows no difference in Kaplan-Meier curves for reaching composite endpoint of death or ICU admission with or without AZ (p = 0.071) NB: HQ but not CQ was associated with decreased propensity adjusted hazard ratio for reaching composite endpoint: HR, 0.68; 95% CI, 0.49-0.95; p = 0.24</td>
</tr>
<tr>
<td>Annie et al. Oktober [49]</td>
<td>Multicentric, retrospective cohort study</td>
<td>Hospitalized patients with confirmed infection</td>
<td>HQ alone (n = 367) vs no HQ (n = 367) HQ + AZ (n = 199) vs no HQ (n = 199) (propensity matched sample taken from 3012 hospitalized patients)</td>
<td>Dosing: not reported</td>
<td>All-cause 30-days mortality</td>
<td>Propensity score matched odds ratio for mortality</td>
</tr>
<tr>
<td>Szente et al. November [50]</td>
<td>Prospective observational study</td>
<td>Outpatients with confirmed infection</td>
<td>HQ, AZ, prednisolone, zinc sulphate, ivermectin and oseltamivir were allowed AZ without HQ or prednisone (n = 106) AZ combined with HQ or prednisone (n = 489) No antiviral treatment (n = 122)</td>
<td>Dosing: HQ: 400mg bid day 1, 400mg od day 2-5 Predni: 1mg/kg od 5 days starting not earlier than day 6 AZ: 500mg od 5 days</td>
<td>Hospitalization risk</td>
<td>Multivariate logistic regression adjusted odds ratio for hospitalization</td>
</tr>
</tbody>
</table>

- Propensity matching was done for HQ and not for AZ
- Factors that may have influenced treatment decisions (e.g. palliative care) or some measures for disease severity (lymphocytes, D-dimers) were not accounted for in multivariate regression
<table>
<thead>
<tr>
<th>Study</th>
<th>Design</th>
<th>Comparator 1</th>
<th>Comparator 2</th>
<th>Outcome</th>
<th>Effect Size</th>
<th>95% CI</th>
<th>p-value</th>
<th>Results</th>
</tr>
</thead>
<tbody>
<tr>
<td>Patel et al.</td>
<td>Systematic review and meta-analysis; uses data Magagnoli et al, Rosenberg et al</td>
<td>HQ + AZ (n = 854) vs SOC (n = 395)</td>
<td>HQ + AZ (n = 854) vs HQ alone (n = 388)</td>
<td>All-cause mortality</td>
<td>Odds ratio for death</td>
<td></td>
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<td>Increased all-cause mortality but causes not assessed</td>
</tr>
<tr>
<td>Das et al.</td>
<td>Meta-analysis using data from Magagnoli et al, Rosenberg et al for AZ assessment</td>
<td>HQ alone (n = 3481)</td>
<td>HQ + AZ (n = 1145)</td>
<td>Neither (n = 1165)</td>
<td>All-cause mortality</td>
<td>Odds ratio for death</td>
<td></td>
<td>- HQ alone vs neither: OR, 0.87; 95% CI, 0.46 – 1.64; p = 0.66</td>
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<td>- HQ + AZ vs neither: OR, 2.84; 95% CI, 2.19 – 3.69; p &lt; 0.001</td>
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<td>- HQ vs HQ + AZ: OR, 0.7; 95% CI, 0.54 – 0.9; p = 0.006</td>
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<td></td>
<td>HQ + AZ associated with increased mortality (HQ alone vs HQ + AZ OR 0.7)</td>
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<td></td>
<td>HQ +/- AZ was associated with increased occurrence of cardiac adverse events but no difference in cardiac adverse events between HQ alone and HQ + AZ</td>
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<td></td>
<td>The outcomes that favoured HQ over HQ + AZ were not cardiac adverse events but mortality rate and development of severe disease; little mechanistic rationale to expect disease worsening with association of AZ and effect may thus be due to residual indication bias</td>
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<tr>
<td>Fiolet et al.</td>
<td>Meta-analysis using data from Ip et al, Magagnoli et al, Rosenberg et al, Ip et al</td>
<td>HQ alone (n = 11932)</td>
<td>AZ + HQ (8081)</td>
<td>Neither (n = 12930)</td>
<td>Mortality</td>
<td>Relative risk for death</td>
<td></td>
<td>- HQ alone vs neither: RR, 0.83; 95% CI, 0.65 – 1.06</td>
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<td>- HQ + AZ vs neither: RR, 1.27; 95% CI, 1.04 – 1.54</td>
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<td>HQ + AZ associated with increased mortality</td>
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<tr>
<td>Yang et al.</td>
<td>Meta-analysis using data from Magagnoli et al, Rosenberg et al</td>
<td>HQ alone (n = 451)</td>
<td>HQ + AZ (n = 854)</td>
<td>vs neither (n = 395)</td>
<td>All-cause mortality, progression to severe illness</td>
<td>Odds ratio for death</td>
<td></td>
<td>- HQ alone vs neither: OR, 1.23; 95% CI, 0.38 – 3.97; p = 0.73</td>
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<td>- HQ + AZ vs neither: OR, 2.34; 95% CI, 1.63 – 3.36; p &lt; 0.0001</td>
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<td>HQ + AZ associated with increased mortality</td>
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<td></td>
<td></td>
<td>Trend towards QT prolongation in HQ treatment did not reach significance</td>
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<td>Duration of follow up (&lt; 14 days or &gt; 14 days) reduces mortality difference (early CV side effects but long term infection reduction?)</td>
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<td>Trend towards increased progression to severe disease in combination treatment; little mechanistic rationale to expect disease worsening with association of AZ and effect may thus be due to residual indication bias</td>
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<tr>
<td>Kashour et al.</td>
<td>Meta-analysis using data from Rosenberg et al, Magagnoli et al, Kuderer et al,</td>
<td>15938 patients to assess effect of HQ</td>
<td>3430 patients to assess effect of HQ + AZ</td>
<td>Short-term mortality</td>
<td>Adjusted OR on short term mortality:</td>
<td></td>
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<td>HQ + AZ associated with increased short-term mortality</td>
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<td>- HQ alone vs neither: effect estimate, 1.05; 95% CI, 0.96 – 1.15; p = 0.647</td>
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<td>- HQ + AZ vs neither: effect estimate, 1.32; 95% CI, 1.00 – 1.75; p = 0.008</td>
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<td>HQ + AZ associated with increased short-term mortality</td>
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<td>Mega et al.</td>
<td>Meta-analysis using data from Magagnoli et al, Rosenberg et al</td>
<td>HQ + AZ (n = 729)</td>
<td>HQ alone (n = 1684)</td>
<td>All-cause mortality, ICU admission, QT prolongation</td>
<td>Odds ratio for composite of death or ICU admission</td>
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<td>- HQ + AZ vs HQ alone: OR, 1.11; 95% CI, 0.54 – 2.28; p = 0.79</td>
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<td>- HQ + AZ vs HQ: OR, 0.88; 95% CI, 0.55 – 1.43; p = 0.61</td>
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AZ: azithromycin, CI: confidence interval, CQ: chloroquine, HQ: hydroxychloroquine, HR: hazard ratio, O2/FiO2: ratio of oxygen saturation (%) over fraction of inspired oxygen (%), MV: mechanical ventilation, OR: odds ratio, RCT: randomized controlled trial, RR: relative risk
References:


is associated with reduced mortality: Findings from the observational multicentre Italian CORIST study. Eur J Intern Med 2020;0. doi:10.1016/j.ejim.2020.08.019


36 Tanriverdi E, ÇÖrtÜk M, Yildirim BiZ, et al. The use of hydroxychloroquine plus azithromycin and early hospital


