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Ambient air pollution and risk of respiratory infection among adults: evidence from the multiethnic study of atherosclerosis (MESA)

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ABSTRACT

Background Air pollution may affect the risk of respiratory infection, though research has focused on uncommon infections or infections in children. Whether ambient air pollutants increase the risk of common acute respiratory infections among adults is uncertain, yet this may help understand whether pollutants influence spread of pandemic respiratory infections like COVID-19.

Objective To estimate the association between ambient air pollutant exposures and respiratory infections in adults.

Methods During five study examinations over 12 years, 6536 participants in the multiethnic study of atherosclerosis (MESA) reported upper respiratory tract infections, bronchitis, pneumonia or febrile illness in the preceding 2 weeks. Using a validated spatiotemporal model, we estimated residential concentrations of ambient $PM_{2.5}$, NO_x and NO_2 for the 2–6 weeks (short-term) and year (long-term) prior to each examination.

Results In this population aged 44–84 years at baseline, 10%–32% of participants reported a recent respiratory infection, depending on month of examination and study region. $PM_{2.5}$, NO_x and NO_2 concentrations over the prior 2–6 weeks were associated with increased reporting of recent respiratory infection, with risk ratios (95% CIs) of 1.04 (1.00 to 1.09), 1.15 (1.10 to 1.20) and 1.21 (1.10 to 1.33), respectively, per increase from 25th to 75th percentile in residential pollutant concentration.

Conclusion Higher short-term exposure to $PM_{2.5}$ and traffic-related pollutants are associated with increased risk of symptomatic acute respiratory infections among adults. These findings may provide an insight into the epidemiology of COVID-19.

Key messages

► Is exposure to ambient air pollutants associated with respiratory infection in adults?

► An increase in residential concentration of $PM_{2.5}$, NO_x and NO_2 from the 25th to the 75th percentile in the previous 2–6 weeks period was associated with 4%, 15% and 21% higher risk of reporting a respiratory infection.

► These findings indicate that fine particulate and traffic-related air pollutants may increase the risk of respiratory infection among adults.

Short-term exposure is associated with worsening of pulmonary function in adults,[10 11] aggravation of asthma and bronchitis in adults and children[12 13] and increased incidence, hospitalisation and mortality due to respiratory tract infections.[14–17]

Short and long-term exposure to air pollutants may impair host immune defenses, resulting in increased susceptibility to respiratory infections.[18 19] Most prior research on air pollutants and respiratory infections has focused on impacts among children[20 21] or individuals with cystic fibrosis.[22–24] However, the link between pollutant exposure and acute respiratory infections among adults has received comparatively less attention, even as mortality rates due to lower respiratory infections (LRI) among children under 5 years have declined globally without corresponding improvements in the adult population.[25]

Acute respiratory infections among adults are common and represent a major source of disease burden.[26] The number of LRI-linked deaths among people above 70 years old rose from nearly 750 000 to 1.1 million annually between 2000 and 2016.[25] Acute respiratory infections account for approximately 5 million emergency department visits in the USA yearly, and for the largest share (41%) of antibiotics prescribed in the adult outpatient

INTRODUCTION

Exposure to ambient air pollutants is a well-established risk factor for several adverse respiratory outcomes.[1–3] Long-term exposure to particulate and gaseous pollutants has been associated with negative impacts on pulmonary function in adults and children,[4 5] increased risk of asthma development,[6] increased severity of chronic obstructive pulmonary disease (COPD) symptoms,[7] faster progression of emphysema[8] and incidence of malignant neoplasms of the lung.[9]

setting.[27 28] Among individuals older than 5 years in 2016, there were approximately 11.6 million LRI episodes in the USA, and about 61 million hospitalisations globally.[25] The ongoing COVID-19 pandemic and ageing global population highlight the urgent need to better understand modifiable risk factors of respiratory infections in the adult population.

While current estimates attribute the majority of the excess burden of disease from ambient air pollutants to cardiovascular causes, a potentially large and poorly estimated fraction of that burden is due to respiratory illnesses.[1] The relative paucity of evidence on the association between air pollution and respiratory infections has recently been noted.[29] We examined the association

Table 1 Baseline characteristics of study participants in the multiethnic study of atherosclerosis (MESA), by report of recent respiratory infection in the first five study examinations*

Characteristic	All participants, n=6536 (100.0%)	Participants who reported an infection at any examination, n=3167 (48.5%)	Participants who reported no infection at all examinations, n=3369 (51.5%)
Age (years), mean (SD)	62.0 (10.2)	61.2 (10.1)	62.8 (10.3)
Race			
White	2516 (38.5)	1157 (36.5)	1359 (40.3)
Chinese-American	787 (12.0)	332 (10.5)	455 (13.5)
African-American	1786 (27.3)	898 (28.4)	888 (26.4)
Hispanic	1447 (22.1)	780 (24.6)	667 (19.8)
Sex			
Female	3445 (52.7)	1765 (55.7)	1680 (49.9)
Male	3091 (47.3)	1402 (44.3)	1689 (50.1)
Education level			
Less than high school	1169 (17.9)	585 (18.5)	584 (17.3)
High school	1187 (18.2)	570 (18.0)	617 (18.3)
Some college or technical certificate	1858 (28.4)	905 (28.6)	953 (28.3)
College or graduate degree	2322 (35.5)	1107 (35.0)	1215 (36.1)
Married/living with a partner	3947 (60.4)	1850 (58.4)	2097 (62.2)
Body mass index, mean (SD)	28.3 (5.5)	28.7 (5.6)	28.0 (5.3)
Ever smoked	3485 (53.3)	1697 (53.6)	1788 (53.1)
Current smoker	741 (11.4)	414 (13.1)	327 (9.8)
Employed	2838 (43.4)	1429 (45.1)	1409 (41.8)
Annual family income			
<$20 000	1597 (24.6)	789 (25.1)	808 (24.2)
$20 000–$39 999	1756 (27.1)	880 (28.0)	876 (26.2)
$40 000–$74 999	1653 (25.5)	816 (26.0)	837 (25.1)
≥$75 000	1474 (22.7)	654 (20.8)	820 (24.5)
Own home or pay mortgage	4413 (67.7)	2045 (64.8)	2368 (70.5)
Excellent self-reported general health status	1465 (46.9)	639 (44.1)	826 (49.3)
No health insurance	464 (7.2)	233 (7.4)	231 (6.9)
On medicaid	616 (9.5)	304 (9.7)	312 (9.3)
Diabetic per ADA 2003 criteria	1052 (16.1)	537 (17.0)	515 (15.3)
Has metabolic syndrome	2485 (38.0)	1235 (39.0)	1250 (37.1)
Moderate and vigorous physical activity, met minutes/week, median (IQR)	3870 (1950–7080)	4065 (2018–7328)	3765 (1905–6840)
Intentional physical activity, met minutes/week, median (IQR)	810 (105–1950)	788 (88–1875)	840 (158–2085)
Self-reported emphysema	73 (1.1)	42 (1.3)	31 (0.9)
Self-reported asthma	253 (3.9)	142 (4.5)	111 (3.3)

*Units are n (%) unless otherwise specified.
ADA, American Diabetic Association.

between long- and short-term exposure to particulate matter less than 2.5 microns in diameter ($PM_{2.5}$) and oxides of nitrogen (NO_x and NO_2) and recent occurrence of upper respiratory tract infections (URTI), bronchitis, pneumonia and febrile illness in a diverse, well-characterised cohort of adults undergoing follow-up in the multiethnic study of atherosclerosis (MESA).

METHODS

MESA is a prospective cohort study of 6814 white, African American, Hispanic and Chinese American individuals recruited from six regions in the USA (Baltimore, MD; New York City, NY; Los Angeles County, CA; Chicago, IL; Winston-Salem, NC and St. Paul, MN) primarily to investigate the progression of subclinical and clinical cardiovascular disease.[30] MESA air is an ancillary study to examine the role of air pollutants in these outcomes.[31] Recruitment and baseline clinical examinations occurred between July 2000 and August 2002. By January 2012, four additional study examinations had been completed. At the first MESA examination in 2000–2002, participants were 44–84 years old and did not have clinical cardiovascular disease. We used data from the first five MESA examinations to quantify the association between long-term and short-term exposure to ambient $PM_{2.5}$, NO_x and NO_2 and the occurrence of respiratory infections in the 2 weeks prior to each study examination. We excluded any observations from study visits at which a participant was taking oral steroids.

Patient and public involvement

The study design, analysis, interpretation of results and writing or editing of this research paper were made without patient involvement.

Exposure estimation

The approach to air pollutant exposure estimation in MESA has been described in detail.[31–33] A hierarchical high-resolution spatiotemporal model is used to predict outdoor concentrations at each participant's residence for each 2 weeks period from January 1999 through the end of participant follow-up. Participants provided residential address information for the entire study duration. The model integrates data from US Environmental Protection Agency Air Quality System regulatory monitors, three types of MESA air-specific monitoring campaigns conducted in the study communities and at homes of study participants and hundreds of geocovariates in a universal kriging and spatial smoothing framework.[33–35] The median (range) cross-validated R^2 values of the prediction models across study regions were 0.79 (0.54–0.85) for $PM_{2.5}$, 0.72 (0.50–0.92) for NO_x, and 0.78 (0.45–0.90) for NO_2. For primary analyses, timescales for assessing associations were selected a priori as the average exposure 1 year before an MESA examination (long-term exposure) and the 4 weeks starting from 14 days before

an examination date (short-term exposure). We also conducted sensitivity analyses using average exposures in the 2-weeks period covering 14–28 days prior to an examination.

Outcome assessment

At each of five study examinations, participants were asked whether they had experienced or been diagnosed with a cold or influenza, febrile illness, sinus infection, bronchitis or pneumonia in the preceding 2 weeks. We constructed a composite binary variable indicating the occurrence of any of the symptoms or diagnoses.

Covariate data

Detailed demographic, anthropometric, lifestyle and medical history data were collected at each MESA clinical examination using self-administered and technician-administered standardised questionnaires.[30] We used data on age (continuous), study region, educational attainment (up to high school or more than high school), sex (male or female), race/ethnicity (white, African American, Hispanic or Chinese American), body mass index in kg/m^2 (<25, 25–30 or ≥30), current smoking status (yes or no), marital status (married or not married), annual family income in US dollars (<20 000; 20 000–39 999; 40 000–74 999; ≥75 000), number of people living in the household, number of children in the household, employment status (employed, retired or unemployed/homemaker), home ownership (outright owner, paying a mortgage and renting or other arrangement), having health insurance (yes or no), diagnosis of diabetes mellitus (based on the 2003 American Diabetic Association fasting criteria algorithm),[36] having metabolic syndrome based on National Cholesterol Education Program guidelines[37] and self-reported doctor-diagnosed asthma or emphysema. Potential confounding by neighbourhood socioeconomic status (NSES) was addressed using an NSES index derived from variables abstracted from the 2000 US Census. As earlier described, this index was obtained by summing z-scores calculated for six NSES variables (median value of owner-occupied housing units, proportion of residents with at least a high school diploma, percentage with at least a bachelor's degree, median household income, percentage with household income <50 000 dollars and percent with professional/managerial occupations).[38]

Statistical analysis

We applied modified Poisson regression with a sandwich linearised estimator of variance to estimate the association between occurrence of respiratory infections in the 2 weeks preceding each of five MESA clinical examinations and average long-term and short-term residential concentrations of ambient $PM_{2.5}$, NO_x and NO_2, adjusting for potential confounders. In contrast to logistic regression, this method for longitudinal binary outcomes[39]
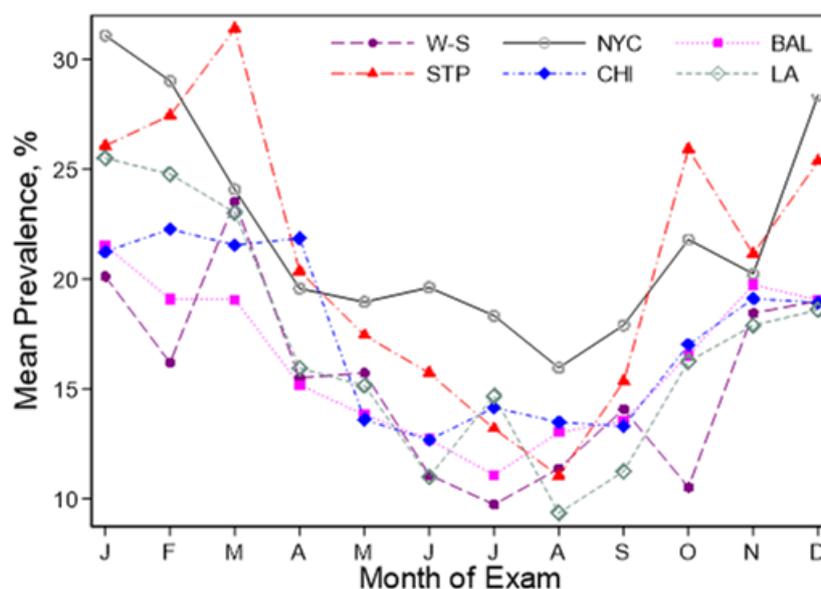
Figure 1 Prevalence of respiratory infection by region and month of examination, across five study examinations. BAL, Baltimore, MD; CHI, Chicago, IL; LA, Los Angeles, CA; NYC, New York City, NY; STP, St Paul, MN; W-S, Winston Salem, NC.

provides a direct measure of the risk ratio and is appropriate when outcomes are common.

We modelled long-term and short-term associations independently, accounting for weather, seasonality and long-term trends. In long-term models, we included smooth functions of calendar time (natural cubic splines with two degrees of freedom per year (df/year)) to account for secular temporal trends. In short-term models, we also used smooth functions of time (splines with six df/year) to account for seasonal and other time-varying influences on infection rates, and additionally included smooth functions of temperature and humidity (four df splines) to deal with potential confounding by weather.

We assessed whether the association between pollutant exposure and recent respiratory infection varied by strata defined by sex, race, current smoking status, age at baseline, diabetes, education level and NSES. Because modification of effects across regions by NSES has been observed in prior MESA studies,[38] we examined models with study region–NSES interactions. Since occurrence of a respiratory infection may be influenced by community-wide levels of infection, we did a secondary analysis to examine if the association between reported infection and ambient pollutant concentration varied by whether the week of the study examination was classified by the Centers for Disease Control (CDC) as meeting the state-level threshold for influenza activity, defined as region-specific periods of two or more consecutive weeks in which each week accounted for more than 2% of the season's total number of specimens that tested positive for influenza in public health laboratories.[40] We also conducted a posthoc assessment of whether level of

pollutant exposure modified the risk of infection during high vs low influenza activity periods.

Sensitivity analyses included estimation of associations separately for each of the five study examinations, and restriction to participants who were present at all examinations. Because febrile symptoms may reflect both respiratory and non-respiratory illness, we conducted an analysis excluding that symptom subcategory. We also evaluated extended models containing self-reported general health status and levels of moderate and vigorous physical activity (metabolic equivalents of minutes/week) for the baseline visit because these covariates were not available at all visits.

RESULTS

After excluding 278 participants who were missing one or more covariates at all MESA examinations, our analytical data set contained 6536 participants at baseline (table 1). We excluded 347 observations from 79 participants, on study visits at which a participant was taking oral steroids. The mean (SD) age at baseline was 62.0 (10.2) years. At baseline, 3445 (52.7%) participants were female, 2516 (38.5%) were white, 741 (11.4%) were current smokers and 2322 (35.5%) had at least a college education. Among Hispanic and African American participants, females and current smokers, a higher fraction reported an infection on at least one study examination than those reporting no infection at all visits (table 1).

By the fifth examination, 4547 participants who met inclusion criteria were still undergoing follow-up, of whom 4197 contributed to the analysis at all five examinations. Baseline characteristics of participants included in the analysis at all visits compared with the 2339 not
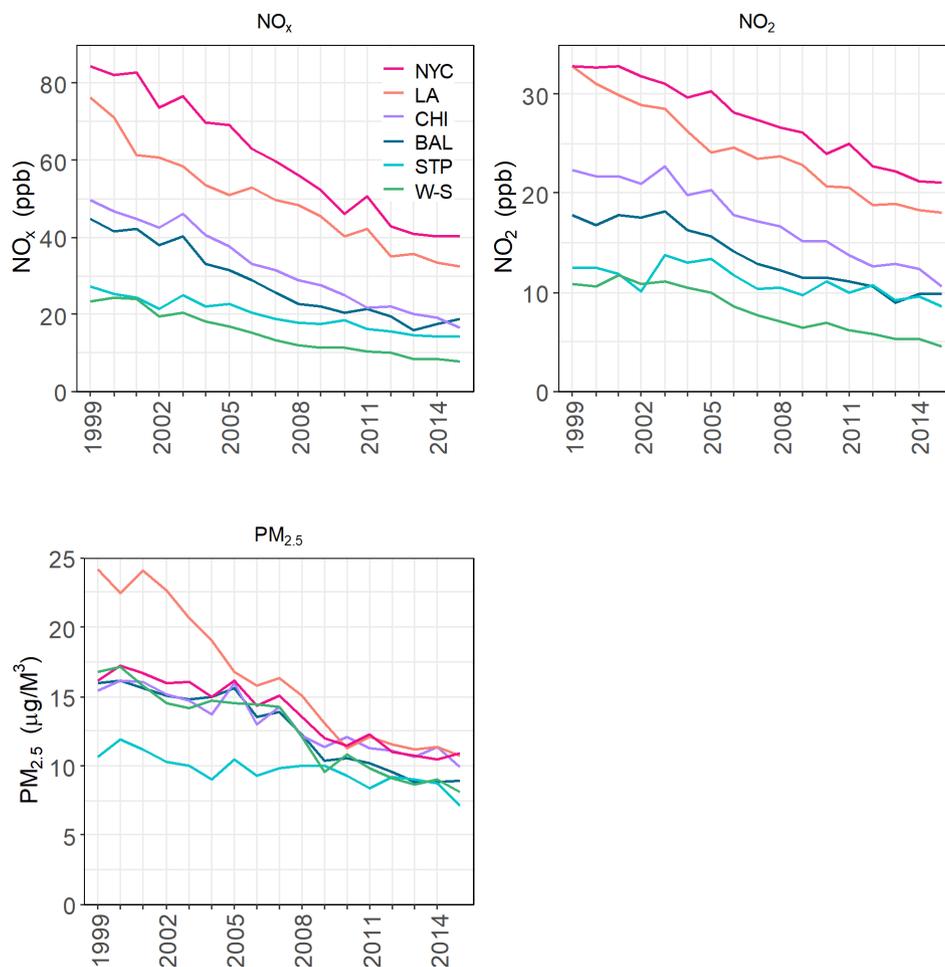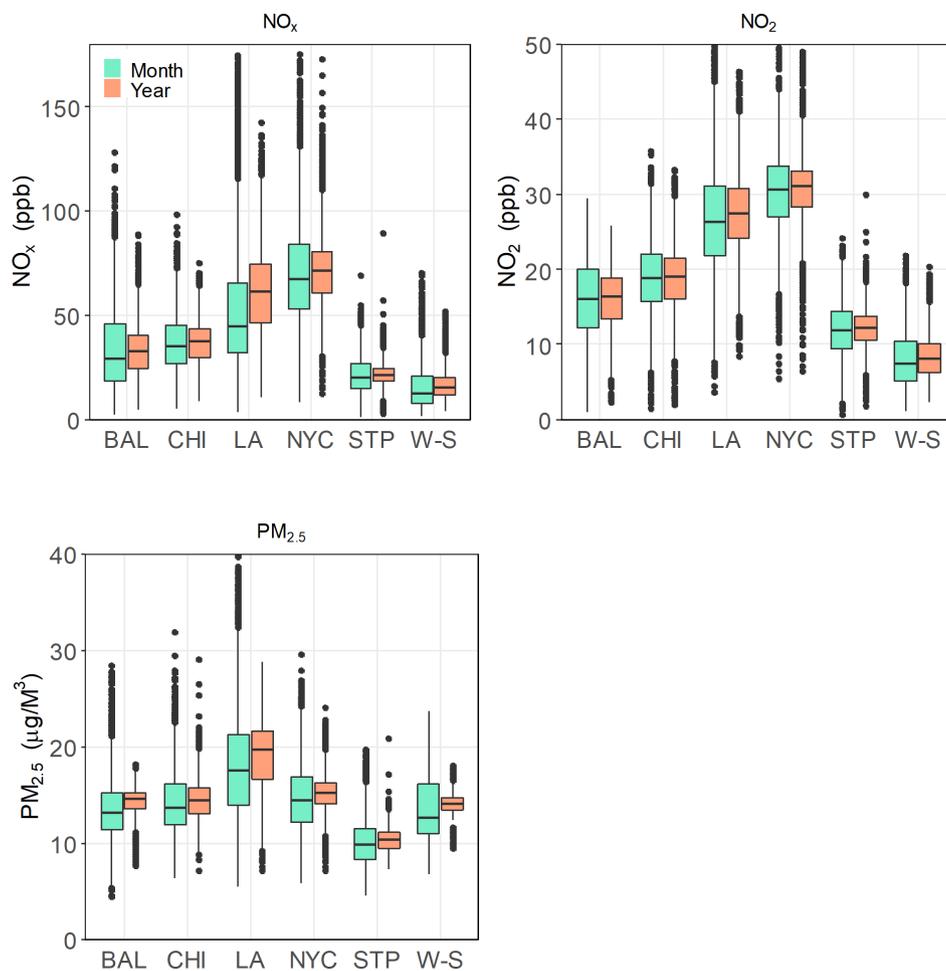
Figure 2 Yearly mean concentrations of PM$_{2.5}$, No$_x$ and NO$_2$ at MESA study participant locations. BAL, Baltimore, MD; CHI, Chicago, IL; LA, Los Angeles, CA; NYC, New York City, NY; STP, St Paul, MN; W-S, Winston Salem, NC. MESA, multiethnic study of atherosclerosis.

included in at least one examination are shown in online supplemental table 1. Generally, participants contributing to the analysis at all examinations were younger and reported relatively higher levels of self-rated health and socioeconomic status. The distribution of study visits per season is shown in online supplemental figure 1; a lower proportion of examination 3 visits occurred in the cooler months.

The proportion of participants reporting a recent respiratory infection at each examination was 20.8%, 18.7%, 18.2%, 17.5% and 13.5%, respectively. The prevalence of infection varied between 10% and 32% depending on study region and month of visit, with lower rates during the warmer months (figure 1). The average concentrations of PM$_{2.5}$, NO$_2$ and NO$_x$ varied by region, but generally declined over the study period (figure 2). There was substantial within-region and between-region variability in pollutant concentrations (figure 3).

In unadjusted models, the risk ratios (95% CIs) for the associations between short-term PM$_{2.5}$, NO$_x$ and NO$_2$ exposure and report of recent respiratory infection were 1.07 (1.03 to 1.13), 1.23 (1.20 to 1.26) and 1.28 (1.17 to 1.39), respectivey. In fully adjusted models,

higher average short-term PM$_{2.5}$, NO$_x$ and NO$_2$ exposure was associated with statistically significantly greater risk of reporting a recent respiratory infection at that visit (PM$_{2.5}$: RR 1.04 (95% CI: 1.00 to 1.09), NO$_x$: 1.15 (1.10 to 1.20) and NO$_2$: 1.21 (1.10 to 1.33) per interquartile range differences of 5.7 µg/m^3, 33.5 ppb and 14.4 ppb, respectively) (figure 4). Results for examination-specific associations were similar, although slightly attenuated and with gradual loss of precision over subsequent examinations (figure 4 and online supplemental table S2). We did not observe statistically significant associations between report of respiratory infection and average long-term exposure (1 year before examination) to any of the pollutants. Results from a short-term exposure sensitivity analysis using average concentrations over a 2-week period before study examination were very similar to those from the 4-week average analysis (online supplemental figure S2) and online supplemental table S2). Additionally, our data did not show evidence of statistically significant modification of the association between short-term exposure to PM$_{2.5}$, NO$_2$ and NO$_x$ and risk of infection by any of the characteristics we examined (online supplemental table S3).

Figure 3 Range of estimated No_x, NO_2, and $PM_{2.5}$ exposures at study participant residential locations over the year (long-term) and the 2–6 weeks (short-term) interval preceding a clinical examination in MESA, by study site. BAL, Baltimore, MD; CHI, Chicago, IL; LA, Los Angeles, CA; NYC, New York City, NY; STP, St Paul, MN; W-S, Winston Salem, NC.MESA, multiethnic study of atherosclerosis.

Quantitative influenza data could only be obtained from one metropolitan area (Los Angeles), and only for the period starting October 2006. Based on these limited data, we did not find clear evidence that the risk

for recent respiratory infection is modified by whether the week of the study visit was classified by the CDC as a influenza week, a designation that denotes a heightened level of circulating influenza-causing pathogens as
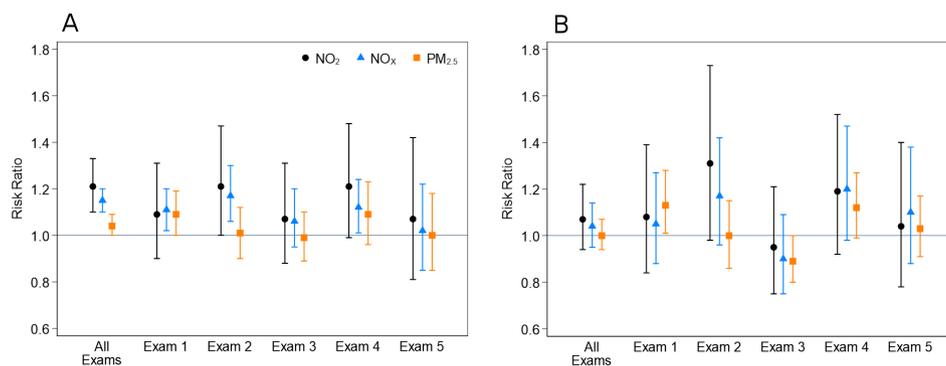
Figure 4 Risk ratios and 95% CIs for associations between respiratory infection and pollutant exposure in the 2–6 weeks (A) and year (B) before interview, for an IQR increase in average pollutant exposure. Results are from model adjusted for study site, race/ethnicity, sex, age, body mass index, marital status, education level, smoking status (ever/never smoked), employment status, annual income, neighbourhood level socioeconomic index, home ownership, number of household members, having health insurance, whether on medicaid, diabetes diagnosis based on ADA criteria and splines of calendar time (and temperature and relative humidity for short-term exposure analysis). ADA, American Diabeteic Association.

tested in public health laboratories across the country (online supplemental table S4). Although our data suggested an increased risk of infection for participants with high average short-term exposure to $PM_{2.5}$ and NO_2 during high-relative to low-influenza activity periods in Los Angeles, the CIs are wide due to the reduced sample size in this subset of examinations (online supplemental table S5). Additionally, there was no effect modification by NSES, based on models that included study region by NSES interaction terms.

Restricting the sample to the 4197 participants who completed all five study visits and were eligible for the study at each visit, we obtained very similar results (online supplemental table S6) and online supplemental figure S3). When we excluded febrile illness, the results remained the same (online supplemental table S7). Models that additionally included physical activity and self-reported health status did not materially alter the results.

DISCUSSION

In a large well-characterised multiethnic cohort observed on up to five occasions over a 12-year period, average exposure to $PM_{2.5}$, NO_x and NO_2 in the 2–6 weeks period preceding a clinic visit was associated with report of a recent respiratory infection. The analysis benefited from a validated exposure assessment method developed for this cohort, taking into account changes in participant address and using study-specific monitoring. The model incorporated both regulatory agency and study-specific monitoring to provide refined spatial and temporal contrasts within and between study regions. These findings suggest that short-term variations in fine particulate and traffic-related air pollution may influence susceptibility to acute respiratory infections among adults.

Comparison with previous studies

Much of the research assessing the association between air pollution and respiratory disease has focused on LRI outcomes among children.[41 42] Reports have also highlighted the link between air pollution and mortality, hospitalisation and emergency department visits, especially due to pneumonia, asthma and exacerbations of COPD, among adults.[43–47] However, the evidence base for the role of air pollutants in the acute manifestation of more common respiratory symptoms and infections among ambulatory adults in the general population is more limited. In a case-crossover study of 57 000 patients in Beijing, $PM_{2.5}$, PM_{10}, NO_2, SO_2 and CO were associated with outpatient visits for URTI, acute bronchitis and community-acquired pneumonia.[48] A retrospective time series analysis in Hong Kong found significant associations between elevated exposures to PM_{10}, ozone and especially NO_2 and general outpatient clinic visits for URTI.[49] Another Chinese time series study of 1.4 million patient visits for respiratory diseases that included pneumonia and acute bronchitis, reported associations with

higher outdoor $PM_{2.5}$, PM_{10}, NO_2 and ozone.[50] In a less-polluted part of China, $PM_{2.5}$, PM_{10} and NO_2 exposures, over the past 7 days were associated with outpatient visits for respiratory disease in a 4 year period.[51]

Our study reflects the context of multiple US cities with comparatively lower pollution levels, and leverages highly resolved exposure information and covariate data collected over a long period in which exposure levels varied substantially.

We observed strong temporal trends in both pollutant concentrations and in the rates of illness reporting among participants. Therefore, we deployed strong statistical controls (adjustment for calendar time and seasonality) which greatly reduced potential for these trends to influence the results. It could be argued that this is overadjustment, if in fact these observed long-term secular trends of reduced reporting of respiratory infections were caused by declines in air pollution. In that case, our estimates of effect may be attenuated relative to the true magnitude of effect.

Acute respiratory infections are major drivers of primary healthcare and antibiotic drug utilisation, suggesting that pollution control may have wide-ranging individual and community health benefits. In a recent analysis, daily rates of telephone calls and e-mails, as well as actual provider visits and emergency department visits related to respiratory symptoms were linked to increased $PM_{2.5}$ levels on that day, with 2.61% more calls and e-mails and 2.63% more provider visits per 100 000 health insurance programme members per day.[52] The same syndromic surveillance approach also demonstrated increased respiratory-related healthcare utilisation for increases in ozone concentrations.[53] An advantage of this design, shared with our approach, is the ability to capture subacute symptoms for which a person may not have otherwise sought medical care.

Evidence pointing to a role for air pollutants in morbidity and mortality due to the novel pandemic coronavirus caused by SARS-CoV-2 continues to emerge, although most reports still drawn from ecological study designs because detailed individual-level data are not yet available.[54–56] Among older adults in the USA, a nationwide study based on county-level COVID-19 death rates suggested that a 1 µg/m³ higher $PM_{2.5}$ exposure was associated with an 11% increase in COVID-19 mortality.[57] In Europe, multiple studies have also provided preliminary indications of positive area-level associations between air pollution ($PM_{2.5}$ and NO_2) and COVID-19 severity, including a report associating an additional 1 µg/m³ of $PM_{2.5}$ at the municipality-level with nine extra COVID-19 cases, three extra hospital admissions and one extra death.[58–60] Initial investigations using limited individual-level data also point to an increase of approximately 7% in mortality rate for each additional 1 µg/m³ in long-term $PM_{2.5}$ exposure.[61] While high-quality representative individual-level data remain scarce, this rapidly evolving body of literature suggests that air pollutants may exacerbate the severity of COVID-19, especially in the presence

of other key risk factors such as advanced age, obesity, and comorbid health conditions.

Possible mechanisms

The mechanisms through which pollutants potentially increase the risk of respiratory infection are still being studied. Particulate and gaseous pollutants are thought to cause respiratory injury through induction of oxidative stress and inflammation in the airways.[18 62] Particles are removed through exhalation, mucociliary clearance or phagocytosis by alveolar macrophages; or they may be translocated through respiratory epithelial surfaces.[9 63] Particle deposition and clearance vary by individual-level factors such as age, sex, smoking status, concomitant conditions like asthma and COPD, activity levels and breathing patterns, which may therefore affect the risk of developing pollutant-related disease.[9] Particles that remain in the respiratory tract can cause inflammation, oxidative stress, inhibition of mucociliary clearance mechanisms, disruption of epithelial tight junctions and impairment of macrophage antimicrobial phagocytic activity, potentially enhancing susceptibility to respiratory infections.[18 64] There is mounting in vitro and in vivo evidence for the pulmonary immunotoxicity of fine particulate matter, although a clear understanding of the underlying mechanisms remains elusive.[65] Particulate matter appears to enhance *Streptococcus pneumoniae* biofilm growth and colonisation of human middle ear epithelium cells in vitro, and to increase pneumococcal colonisation of the nasopharynx in an animal model.[66]

Oxides of nitrogen are absorbed in both the upper respiratory tract and pulmonary regions, and their reactions with surface epithelial fluids may result in oxidative stress, while absorbed components may cause deep lung irritation and immunomodulation of host response to infection, possibly increasing susceptibility or worsening the severity of existing infection.[67] In the Swiss Cohort Study on Air Pollution and Lung and Heart Disease in Adults (SAPALDIA), increased exposure to NO_2 was linked to higher risk of respiratory problems leading to doctor visits among participants with underlying COPD and chronic bronchitis.[68] This suggests that particulate and traffic-related pollutants may elevate the virulence of secondary pathogenic agents in the event of concurrent exposure in the human respiratory tract, partly through inhibition of immune function. Some investigators have hypothesised that air pollution may increase the mortality risks associated with the novel epidemic Severe Acute Respiratory Syndrome Coronavirus-2.[69]

Strengths and limitations

This study has some limitations. Sufficiently detailed CDC influenza activity data for our period of interest were not readily available for most study regions, reducing our ability to assess the interesting potential interactive role of community-level influenza activity on the pollutant-illness relationship. The outcomes were self-reported,

although trained research staff conducted the interviews consistently over follow-up. In addition, we used outdoor residential pollutant concentrations that do not take into account the actual microenvironments in which participants spend time. Importantly, we utilised high-quality exposure predictions in an ethnically and geographically diverse and well-characterised cohort that permitted ample confounder control and exploitation of fine spatiotemporal variations in exposure.

CONCLUSION AND RECOMMENDATIONS

Our study provides evidence that higher exposure to $PM_{2.5}$, NO_x and NO_2 contributes to an increased risk of common but mostly self-limited respiratory infections. The potential burden of these illnesses—in lost productivity, healthcare utilisation and antibiotic use—suggests that control of particulate and traffic-related pollution may provide substantial individual and societal benefit. In light of the increased attention to respiratory infections in the time of the COVID-19 pandemic, these results also may provide an insight into the potential role of pollutants in modifying the risk of novel viral infections. Given the known deleterious effects of pollutants on the respiratory system, including increasing the risk of infections, and preliminary linkages to COVID-19 morbidity and mortality, it is suggested that reducing pollutant exposure may prevent severe illness and death due to COVID-19. Because respiratory infections, possibly including COVID-19, will remain major health problems among adults, two recommendations can be inferred from our study. First, it may be beneficial for older adults who spend substantial amounts of time in the home to utilise air purification equipment such as filters. Second, multipronged regulatory actions to control pollution at the societal level remain critical public health priorities including science-based adjusments of air quality standards and concerted efforts and resources in areas with high concentrations of particulate and traffic-related pollutants.

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Ambient air pollution and risk of respiratory infection among adults: Evidence from the Multi-Ethnic Study of Atherosclerosis

Kipruto Kirwa, Carly M Eckert, Sverre Vedal, Anjum Hajat, Joel D Kaufman

Online Supplementary Data

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Table S1. Baseline characteristics of study participants followed-up in the Multi-Ethnic Study of Atherosclerosis (MESA), by contribution to all study visits.*

Characteristic	Participants contributing to analysis at all visits *n* = 4,197	Participants who dropped out during follow-up or were ineligible for analysis on at least one examination *n* = 2,339
Age (years), mean (SD)	60.3 (9.6)	65.1 (10.7)
Study region		
Winston-Salem	698 (16.6)	318 (13.6)
New York City	743 (17.7)	323 (13.8)
Baltimore	567 (13.5)	455 (19.5)
St. Paul	688 (16.4)	345 (14.7)
Chicago	812 (19.3)	318 (13.6)
Los Angeles	689 (16.4)	580 (24.8)
Race		
White	1,717 (40.9)	799 (34.2)
Chinese-American	498 (11.9)	289 (12.4)
African-American	1,099 (26.2)	687 (29.4)
Hispanic	883 (21.0)	564 (24.1)
Sex		
Female	2,197 (52.3)	1,248 (53.4)
Male	2,000 (47.7)	1,091 (46.6)
Education level		

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Less than high school	579 (13.8)	590 (25.2)
High school	725 (17.3)	462 (19.8)
Some college or technical certificate	1,215 (28.9)	643 (27.5)
College or graduate degree	1,678 (40.0)	644 (27.5)
Married/living with a partner	2,644 (63.0)	1,303 (55.7)
Body mass index, mean (SD)	28.3 (5.3)	28.4 (5.7)
Ever smoked	2,222 (52.9)	1,263 (54.0)
Current smoker	420 (10.0)	321 (13.9)
Employed	2,055 (49.0)	783 (33.5)
Annual family income, US dollars		
<20,000	801 (19.2)	796 (34.4)
20,000-39,999	1,083 (26.0)	673 (29.1)
40,000-74,999	1,150 (27.6)	503 (21.8)
≥75,000	1,134 (27.2)	340 (14.7)
Own home or pay mortgage (not rent)	2,968 (70.9)	1,445 (62.1)
Excellent self-reported general health status	906 (53.5)	559 (39.0)
No health insurance	241 (5.8)	223 (9.7)
On Medicaid	322 (7.7)	294 (12.8)
Diabetic per ADA 2003 criteria	599 (14.3)	453 (19.4)
Has metabolic syndrome	1,504 (35.9)	981 (42.0)

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Moderate and vigorous physical activity, Met Mins/Week, median (IQR)	4,219 (2,160–7,410)	3,330 (1,508–6,375)
Intentional physical activity, Met Mins/Week, median (IQR)	840 (210–2,062)	712 (0–1,800)
Self-reported emphysema	23 (0.5)	50 (2.1)
Self-reported asthma	136 (3.2)	117 (5.0)

* Units are *n* (%) unless otherwise specified. ADA – American Diabetic Association, IQR – Interquartile range, SD – Standard Deviation

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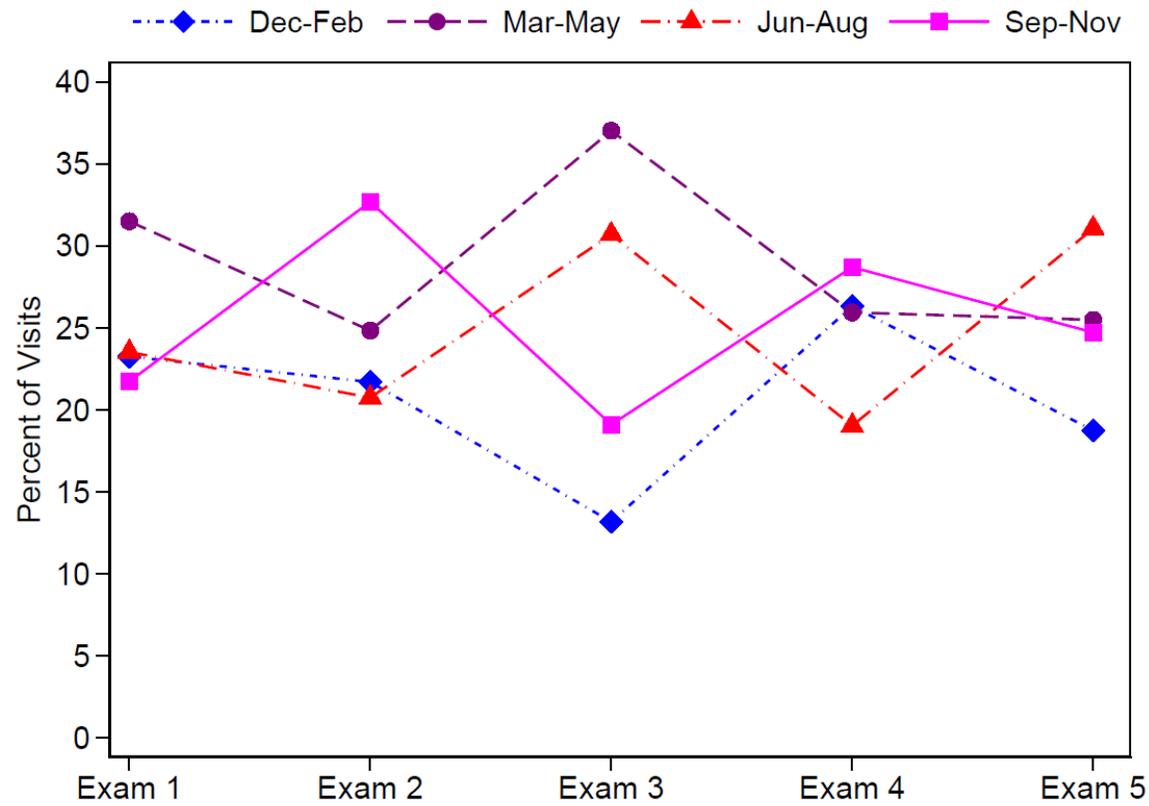
Figure S1. Distribution of study examinations per season in MESA.

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Table S2. Risk ratios (95% CIs) for associations between respiratory infection and average exposure to $PM_{2.5}$, NO_x, and NO_2 among MESA participants, for entire study duration and per study exam.

Pollutant	Examination	Annual average		28-day average		14-day average	
		RR (95% CI)	*p*-value	RR (95% CI)	*p*-value	RR (95% CI)	*p*-value
$PM_{2.5}$	Overall (1–5)	1.00 (0.94, 1.07)	0.980	1.04 (1.00, 1.09)	0.078	1.03 (0.99, 1.08)	0.102
	1	1.13 (1.01, 1.28)	0.038	1.09 (0.99, 1.19)	0.075	1.05 (0.97, 1.14)	0.220
	2	1.00 (0.86, 1.15)	0.957	1.01 (0.90, 1.12)	0.926	1.00 (0.92, 1.10)	0.925
	3	0.89 (0.80, 1.00)	0.047	0.99 (0.89, 1.10)	0.878	0.98 (0.89, 1.07)	0.620
	4	1.12 (0.99, 1.27)	0.074	1.09 (0.96, 1.23)	0.195	1.10 (0.99, 1.24)	0.087
	5	1.03 (0.91, 1.17)	0.666	1.00 (0.85, 1.18)	0.997	1.00 (0.86, 1.16)	0.979
NO_x	Overall (1–5)	1.04 (0.95, 1.14)	0.405	1.15 (1.10, 1.20)	0.000	1.12 (1.08, 1.16)	0.000
	1	1.05 (0.88, 1.27)	0.578	1.11 (1.02, 1.20)	0.012	1.08 (1.00, 1.17)	0.046
	2	1.17 (0.96, 1.42)	0.126	1.17 (1.06, 1.30)	0.003	1.12 (1.02, 1.23)	0.022
	3	0.90 (0.75, 1.09)	0.278	1.06 (0.94, 1.20)	0.314	1.05 (0.95, 1.17)	0.338
	4	1.20 (0.98, 1.47)	0.076	1.12 (1.00, 1.24)	0.048	1.10 (0.99, 1.22)	0.073
	5	1.10 (0.88, 1.38)	0.404	1.02 (0.85, 1.22)	0.843	1.02 (0.86, 1.21)	0.816
NO_2	Overall (1–5)	1.07 (0.94, 1.22)	0.301	1.21 (1.10, 1.33)	0.000	1.15 (1.05, 1.25)	0.002
	1	1.08 (0.84, 1.39)	0.567	1.09 (0.90, 1.31)	0.394	1.03 (0.87, 1.22)	0.761
	2	1.31 (0.98, 1.73)	0.064	1.21 (1.00, 1.47)	0.051	1.12 (0.93, 1.34)	0.226
	3	0.95 (0.75, 1.21)	0.694	1.07 (0.88, 1.31)	0.490	1.04 (0.87, 1.25)	0.650
	4	1.19 (0.92, 1.52)	0.180	1.21 (0.99, 1.48)	0.066	1.15 (0.95, 1.40)	0.154
	5	1.04 (0.78, 1.40)	0.769	1.07 (0.81, 1.42)	0.610	1.07 (0.82, 1.39)	0.626
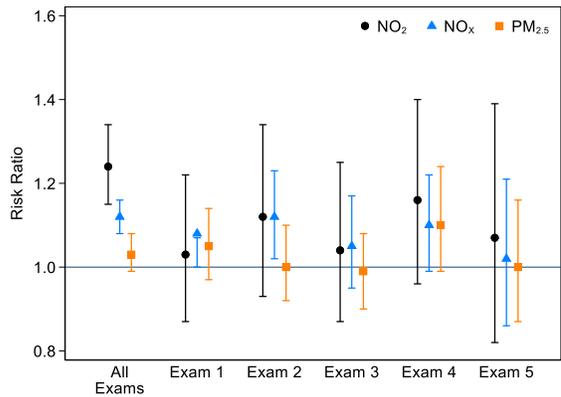
Figure S2. Risk ratios and 95% confidence intervals for associations between respiratory infection and pollutant exposure in the two to four week period before a study examination, for an interquartile range increase in average pollutant exposure. Results are from a model adjusted for study site, race/ethnicity, sex, age, body mass index, marital status, education level, smoking status (ever/never smoked), employment status, annual income, neighborhood level socio-economic index, home ownership, number of household members, having health insurance, whether on Medicaid, diabetes diagnosis based on ADA criteria and natural cubic splines of temperature and relative humidity (four degrees of freedom).

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Table S3. Risk ratios (95% CIs) of respiratory infection for an interquartile range increase in average air pollutant exposure in the two-to-six week period before interview among participants in the MESA study, stratified by select characteristics.

Characteristic	$PM_{2.5}$		NO_x		NO_2	
	Risk ratio (95% CI)	p-value	Risk ratio (95% CI)	p-value	Risk ratio (95% CI)	p-value
Age (years)		0.522		0.428		0.494
Less than 65	1.03 (0.97, 1.10)		1.13 (1.06, 1.20)		1.16 (1.02, 1.32)	
65 or more	1.05 (0.99, 1.12)		1.18 (1.11, 1.24)		1.26 (1.11, 1.44)	
Race/ethnicity		0.853		0.141		0.431
White	1.01 (0.93, 1.10)		1.19 (1.08, 1.32)		1.14 (0.98, 1.33)	
Chinese American	1.16 (0.94, 1.43)		1.25 (1.07, 1.46)		1.37 (1.04, 1.80)	
African American	1.01 (0.94, 1.08)		1.10 (1.01, 1.20)		1.16 (0.98, 1.38)	
Hispanic	1.06 (0.97, 1.17)		1.14 (1.04, 1.24)		1.13 (0.93, 1.36)	
Sex		0.148		0.688		0.861
Female	1.05 (0.99, 1.11)		1.13 (1.07, 1.19)		1.16 (1.03, 1.31)	
Male	1.03 (0.95, 1.10)		1.18 (1.10, 1.26)		1.26 (1.09, 1.46)	
Ever smoked		0.494		0.893		0.370
No	1.06 (0.99, 1.13)		1.16 (1.09, 1.23)		1.16 (1.01, 1.33)	
Yes	1.03 (0.97, 1.09)		1.14 (1.08, 1.21)		1.25 (1.10, 1.42)	
Diabetic		0.137		0.133		0.283
No	1.05 (1.00, 1.11)		1.17 (1.12, 1.23)		1.23 (1.11, 1.36)	
Yes	1.00 (0.90, 1.11)		1.07 (0.96, 1.19)		1.14 (0.89, 1.45)	
Metabolic syndrome		0.161		0.279		0.695
No	1.07 (1.01, 1.13)		1.18 (1.12, 1.24)		1.21 (1.08, 1.36)	
Yes	1.01 (0.94, 1.08)		1.12 (1.05, 1.19)		1.19 (1.03, 1.38)	
Neighborhood SES index at baseline		0.997		0.237		0.378
≤0	1.02 (0.94, 1.10)		1.16 (1.06, 1.27)		1.15 (0.99, 1.32)	
>0	1.05 (0.99, 1.11)		1.15 (1.10, 1.21)		1.25 (1.11, 1.42)	
Education level		0.715		0.212		0.843
Less than high school	1.06 (0.96, 1.16)		1.10 (1.01, 1.20)		1.04 (0.87, 1.26)	
High school	1.02 (0.92, 1.14)		1.23 (1.11, 1.36)		1.23 (0.96, 1.56)	
Some college	1.03 (0.94, 1.12)		1.12 (1.04, 1.21)		1.23 (1.03, 1.47)	
College or graduate	1.05 (0.97, 1.14)		1.21 (1.11, 1.32)		1.27 (1.09, 1.47)	
Household income, US dollars		0.418		0.175		0.419
< 20,000	1.06 (0.97, 1.15)		1.10 (1.03, 1.18)		1.12 (0.95, 1.32)	
20,000 – 39,999	1.08 (0.98, 1.17)		1.20 (1.10, 1.31)		1.19 (0.97, 1.46)	
40,000 – 74,999	0.96 (0.88, 1.05)		1.14 (1.04, 1.26)		1.22 (1.02, 1.46)	
≥ 75,000	1.05 (0.94, 1.16)		1.19 (1.06, 1.34)		1.18 (0.99, 1.42)	

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Table S4. Risk ratios and 95% CIs of respiratory infection for an interquartile range increase in pollutant exposure, stratified by influenza activity-week threshold, based on data from the Los Angeles study area*

Characteristic	Number of participants	Number of visits	Risk ratio (95% CI)	*p*-value
Long-term (1-year) exposure				
$PM_{2.5}$				0.412
Flu week	607	702	1.03 (0.83, 1.28)	
Non-flu week	607	702	1.13 (0.49, 2.62)	
NO_x				0.479
Flu week	607	702	1.11 (0.90, 1.38)	
Non-flu week	607	702	1.26 (0.79, 2.02)	
NO_2				0.472
Flu week	607	702	1.06 (0.88, 1.28)	
Non-flu week	607	702	1.12 (0.80, 1.57)	
Short-term (4-week) exposure				
$PM_{2.5}$				0.088
Flu week	608	703	1.01 (0.93, 1.10)	
Non-flu week	608	703	1.05 (0.78, 1.41)	
NO_x				0.137
Flu week	608	703	1.08 (0.94, 1.24)	
Non-flu week	608	703	1.11 (0.92, 1.35)	
NO_2				0.161
Flu week	608	703	1.09 (0.95, 1.25)	
Non-flu week	608	703	1.18 (0.90, 1.55)	
Short-term (2-week) exposure				
$PM_{2.5}$				0.135
Flu week	608	703	1.01 (0.88, 1.15)	
Non-flu week	608	703	1.02 (0.80, 1.30)	
NO_x				0.161
Flu week	608	703	1.06 (0.95, 1.19)	
Non-flu week	608	703	1.10 (0.91, 1.33)	
NO_2				0.149
Flu week	608	703	1.07 (0.94, 1.22)	
Non-flu week	608	703	1.15 (0.88, 1.49)	

* Flu week is defined as periods of two or more consecutive weeks in which each week accounted
for more than 2% of the season's total number of specimens that tested positive for influenza in
public health laboratories.

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Table S5. Risk ratios and 95% CIs of respiratory infection during high versus low influenza
activity weeks, stratified by exposure to lower or higher than median pollutant levels in the two-
to-six weeks prior to study examination, based on data from the Los Angeles study area*

Characteristic	Risk ratio (95% CI)	p-value
$PM_{2.5}$		0.263
Above median	3.76 (1.56, 9.10)	
Below median	0.91 (0.15, 5.57)	
NO_2		0.298
Above median	3.31 (1.39, 7.89)	
Below median	0.79 (0.12, 5.28)	

* Median pollutant concentrations for the period when influenza activity data were available
were 11.4 μg/m^3 for $PM_{2.5}$ and 24.9 ppb for NO_2 in Los Angeles. Model convergence was not
achieved for NOx. Periods meeting influenza activity thresholds are defined by the Centers for
Disease Control as two or more consecutive weeks in which each week accounted for more than
2% of the season's total number of specimens that tested positive for influenza in public health
laboratories. P-value corresponds to a joint Wald test of equality of coefficients across strata of
variable.

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Table S6. Risk ratios (95% CIs) for associations between respiratory infection and average exposure to $PM_{2.5}$, NO_x, and NO_2 among 4,197 MESA participants who were evaluated at all five study examinations, for entire study duration and per study exam.

Pollutant	Examination	Annual average		28-day average		14-day average	
		RR (95% CI)	*p*-value	RR (95% CI)	*p*-value	RR (95% CI)	*p*-value
$PM_{2.5}$	Overall (1–5)	0.99 (0.91, 1.08)	0.844	1.03 (0.97, 1.09)	0.342	1.03 (0.98, 1.09)	0.190
	1	1.12 (0.98, 1.28)	0.088	1.07 (0.94, 1.21)	0.300	1.07 (0.96, 1.19)	0.240
	2	0.94 (0.80, 1.11)	0.486	0.98 (0.86, 1.12)	0.803	0.99 (0.89, 1.11)	0.865
	3	0.92 (0.81, 1.05)	0.226	1.01 (0.89, 1.15)	0.862	1.00 (0.89, 1.12)	0.989
	4	1.17 (1.02, 1.34)	0.029	1.09 (0.95, 1.26)	0.229	1.13 (1.00, 1.29)	0.054
	5	1.01 (0.89, 1.15)	0.860	0.97 (0.82, 1.15)	0.753	0.97 (0.83, 1.13)	0.660
NO_x	Overall (1–5)	1.03 (0.93, 1.15)	0.555	1.13 (1.08, 1.20)	0.000	1.11 (1.05, 1.16)	0.000
	1	1.10 (0.87, 1.38)	0.426	1.11 (0.99, 1.24)	0.073	1.12 (1.02, 1.24)	0.024
	2	1.11 (0.86, 1.43)	0.414	1.14 (1.00, 1.30)	0.053	1.06 (0.93, 1.19)	0.396
	3	0.93 (0.74, 1.16)	0.506	1.07 (0.93, 1.24)	0.354	1.03 (0.91, 1.18)	0.618
	4	1.27 (1.02, 1.58)	0.035	1.10 (0.97, 1.25)	0.143	1.08 (0.95, 1.22)	0.222
	5	1.08 (0.85, 1.37)	0.515	1.03 (0.85, 1.24)	0.780	1.02 (0.85, 1.22)	0.847
NO_2	Overall (1–5)	1.08 (0.93, 1.26)	0.319	1.20 (1.07, 1.34)	0.002	1.14 (1.02, 1.26)	0.016
	1	1.07 (0.78, 1.47)	0.671	1.05 (0.82, 1.34)	0.698	1.09 (0.87, 1.35)	0.466
	2	1.30 (0.91, 1.86)	0.154	1.20 (0.93, 1.54)	0.155	1.06 (0.84, 1.33)	0.645
	3	1.02 (0.77, 1.36)	0.888	1.12 (0.88, 1.42)	0.360	1.05 (0.84, 1.30)	0.679
	4	1.34 (1.02, 1.76)	0.034	1.28 (1.02, 1.60)	0.030	1.20 (0.97, 1.48)	0.091
	5	1.01 (0.74, 1.37)	0.961	1.07 (0.80, 1.43)	0.658	1.04 (0.78, 1.37)	0.798

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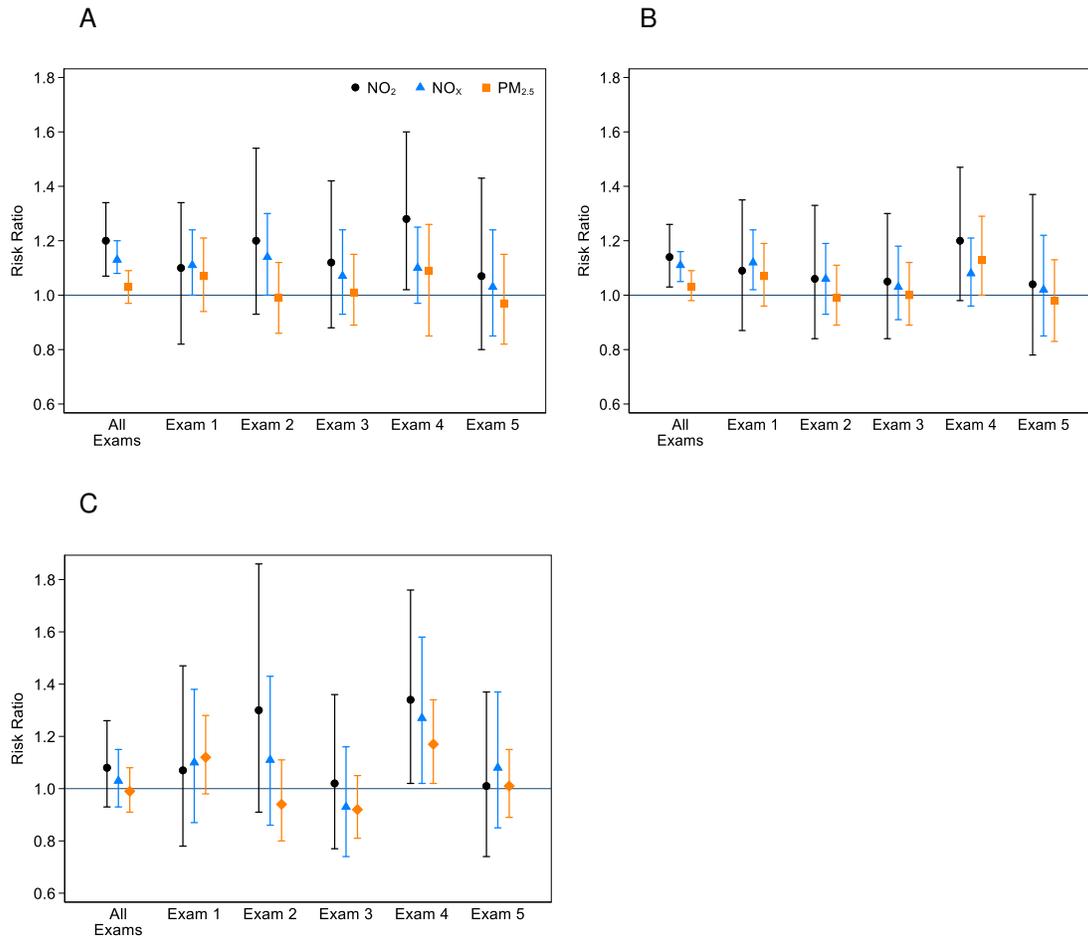
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Figure S3. Risk ratios and 95% confidence intervals for associations between respiratory infection and pollutant exposure among 4,197 MESA participants who were evaluated at all five study examinations, in the two-to-six weeks (A), two-to-four weeks (B), and year (C) prior to study visit, for an interquartile range increase in average pollutant exposure.

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Table S7. Risk ratios and 95% confidence intervals for associations between respiratory
infection and pollutant exposure in the two-to-six weeks before study examination when
excluding report of febrile illness from the outcome definition, for an interquartile range
increase in average pollutant exposure.

Pollutant	Risk ratio (95% CI)	
	Main result in Figure 4	Excluding febrile illness
$PM_{2.5}$	1.04 (1.00, 1.09)	1.04 (0.99, 1.09)
NO_x	1.15 (1.10, 1.20)	1.15 (1.10, 1.20)
NO_2	1.21 (1.10, 1.33)	1.21 (1.10, 1.33)